

Bayesian, Fiducial, and Frequentist (BFF): Best Friends Forever?

Xiao-Li Meng

Department of Statistics, Harvard University

- Liu & Meng (2106) **There Is Individualized Treatment. Why Not Individualized Inference?** *Annual Review of Statistics and Its Application*, 3: 79-111
- Liu & Meng (2014). **A Fruitful Resolution To Simpson's Paradox via Multi-Resolution Inference.** *The American Statistician*, 68: 17-29.
- Meng (2014). **A Trio of Inference Problems That Could Win You a Nobel Prize in Statistics (if you help fund it).** *In the Past, Present, and Future of Statistical Science (Eds: X. Lin, et. al.)*, 535-560.

What is *inference*? Katie's answer ...



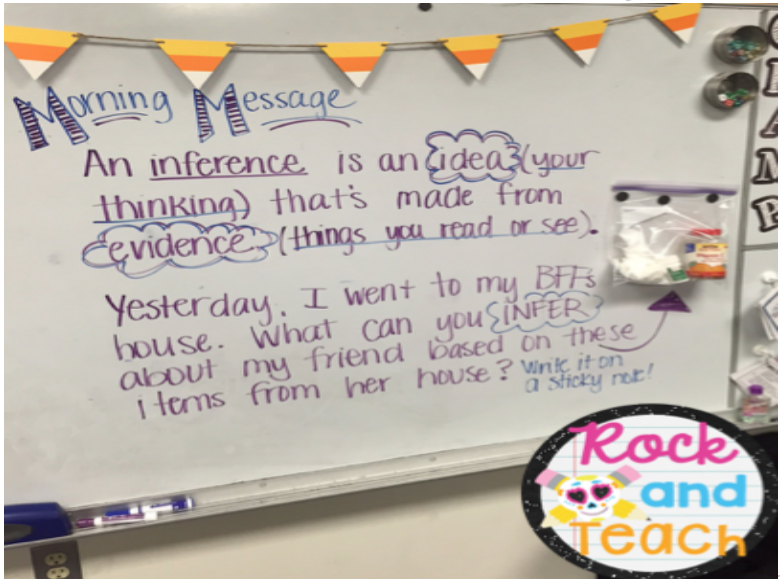
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Summary



But what is *Statistical/Probabilistic* Inference?

- An ultimate intellectual game: **“to guess wisely and to guess meaningfully the errors in our guesses.”**
(*XL-Files*, Oct 2015)

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Summary

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Pure Frequentist (Fully unconditional)

Most Robust but Least Relevant

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Most Robust but Least Relevant

Pure Bayesian (Fully conditional)

Most Relevant but Least Robust

But life is about *compromise*:

Conditional frequentist, Objective Bayesian, Fiducial ...



It all depends on which *Replications* you want ...

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Summary

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Summary

Statistical Model via Stochastic Representation

$$\underbrace{D}_{\text{Data}} = G(\underbrace{\theta}_{\text{Signal}}, \underbrace{U}_{\text{Noise}}) \quad (S)$$

It all depends on which *Replications* you want ...

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Statistical Model via Stochastic Representation

$$\underbrace{D}_{\text{Data}} = G(\underbrace{\theta}_{\text{Signal}}, \underbrace{U}_{\text{Noise}}) \quad (S)$$

Ex: $D = \{X_1, \dots, X_n\}$, where

$$X_i = \theta + U_i, \quad U_i \stackrel{\text{iid}}{\sim} N(0, 1),$$

and $U = \{U_i, i = 1, \dots, n\}$ represents “God’s Uncertainty”

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- Frequentist: Fix parameter θ , vary D

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- Frequentist: Fix parameter θ , vary D
- Bayesian: Fix data D , vary θ

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- Frequentist: Fix parameter θ , vary D
- Bayesian: Fix data D , vary θ
- Fiducial: Fix neither, but vary U , subject to the constraint (S) (or implied constraints with $A(U)$ fixed)

The differences are in the replications ...

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Summary

θ
|
 D

The differences are in the replications ...

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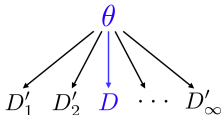
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Summary

θ
|
 D

Frequentist Inference



$$p(D'|\theta)$$

The differences are in the replications ...

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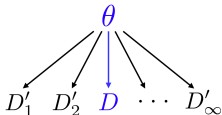
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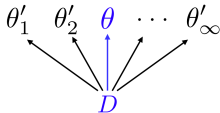
θ
|
 D

Frequentist Inference



$$p(D'|\theta)$$

Bayesian Inference



$$p(\theta'|D)$$

$$\propto p(D|\theta')\pi_0(\theta')$$

The differences are in the replications ...

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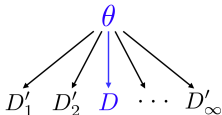
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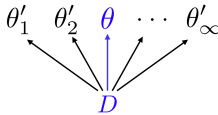
θ
↓
 D

Frequentist Inference



$$p(D'|\theta)$$

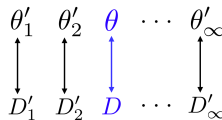
Bayesian Inference



$$p(\theta'|D)$$

$$\propto p(D|\theta')\pi_0(\theta')$$

Fiducial Inference



$$p(D', \theta'|A(U))$$

$$= p(D'|\theta', A(U))\pi(\theta')$$



Illustrate BFF for $X \sim N(\theta, 1)$

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Summary

Frequentist



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Frequentist

Bayesian



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Summary

Frequentist

Bayesian

Fiducial



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Summary

Frequentist

Bayesian

Fiducial

Sampling Dist.

$X|\theta \sim N(\theta, 1)$

Illustrate BFF for $X \sim N(\theta, 1)$

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Summary

Frequentist

Sampling Dist.
 $X|\theta \sim N(\theta, 1)$

Bayesian

+ Prior Dist.
 $\pi_0(\theta) \propto 1$

Fiducial

Illustrate BFF for $X \sim N(\theta, 1)$

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Fiducial

God's U Dist.

$$X - \theta = U \sim N(0, 1)$$

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Confidence Interval

$$(X - z_p, X + z_p)$$

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Posterior Dist.

$$\theta|X \sim N(X, 1)$$

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Fiducial Dist.

$$\theta = X + U \sim N(X, 1)$$

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Confidence Dist.

$$N(X, 1)$$

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Sampling Dist.

$$X|\theta \sim N(\theta, 1)$$

Confidence Dist.

$$N(X, 1)$$

Confidence Interval

$$(X - z_p, X + z_p)$$

Generate for $i = 1, \dots$

$X_i|\theta \sim N(\theta, 1)$, then
 $(X_i - 1.96, X_i + 1.96)$
 covers θ 95% of times

Bayesian

+ Prior Dist.

$$\pi_0(\theta) \propto 1$$

Posterior Dist.

$$\theta|X \sim N(X, 1)$$

Posterior Interval

$$(X - z_p, X + z_p)$$

Fiducial

God's U Dist.

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Fiducial Dist.

$$\theta = X + U \sim N(X, 1)$$

Fiducial Interval

$$(X - z_p, X + z_p)$$

Illustrate BFF for $X \sim N(\theta, 1)$

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Confidence Dist.

$$N(X, 1)$$

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Bayesian

+ Prior Dist.

$$\pi_0(\theta) \propto 1$$

Posterior Dist.

$$\theta|X \sim N(X, 1)$$

Posterior Interval

$$(X - z_p, X + z_p)$$

Generate for $i = 1, \dots$

$\theta_i|X \sim N(X, 1)$, then
 $(X - 1.96, X + 1.96)$
covers θ_i 95% of times

Fiducial

God's U Dist.

$$X - \theta = U \sim N(0, 1)$$

Fiducial Dist.

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Fiducial Interval

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Posterior Dist.

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Generate for $i = 1, \dots$

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covers θ_i 95% of times

Fiducial

God's U Dist.

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Fiducial Dist.

$$\theta = X + U \sim N(X, 1)$$

Fiducial Interval

$$(X - z_p, X + z_p)$$

Generate for $i = 1, \dots$

$\theta_i \sim \text{any } \pi(\theta)$, &
 $X_i|\theta_i \sim N(\theta_i, 1)$, then
 $(X_i - 1.96, X_i + 1.96)$
covers θ_i 95% of times



Finding the Right “Control Population”: Treating Data as Your Patient

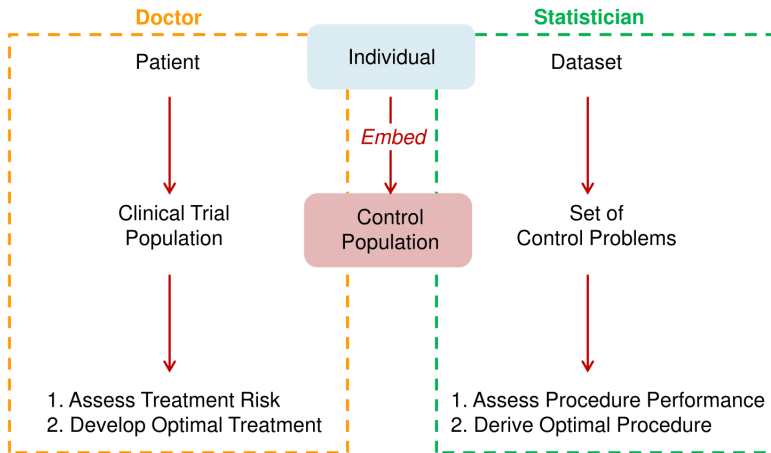
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Summary





The Inevitable Statistical “Bootstrap”: Creating Internal Replications

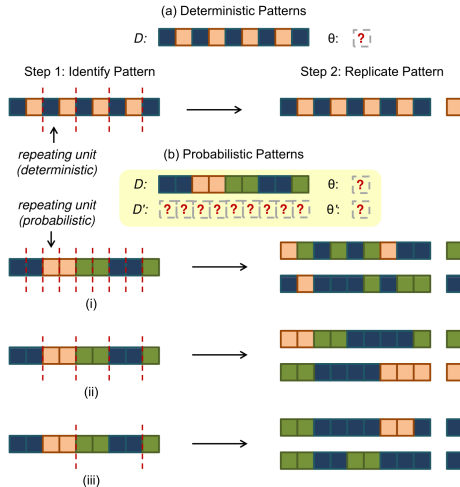
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Summary



Relevant Controls/Replications are always needed

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Summary

Method Type	Error on Actual Problem Δ	Average Error over Relevant Controls $\bar{\Delta}'$	References
Point Estimate <i>Goal:</i> Give our best guess, $\hat{\theta}$, for value of θ .	$L(\theta, \hat{\theta})$ <i>Loss:</i> Specify how "far" θ is from $\hat{\theta}$ via loss function.	<i>Risk:</i> The average loss of an estimator over control problems (D', θ') .	Robinson 1979b Rukhin 1988, Lu and Berger 1989, Fourdrinier and Wells 2012
Set Estimate <i>Goal:</i> Identify set, $C(D)$, of likely values for θ .	$I(\theta \notin C(D))$ * <i>Coverage:</i> Does our set contain the true value of θ ?	<i>Non-Coverage Probability:</i> Proportion of times a set estimate, e.g. interval estimate, fails to contain the true value of θ' .	Casella 1992, Goutis and Casella 1995, Robinson 1979a, Berger 1988
Hypothesis Test <i>Goal:</i> Should we reject a null hypothesis, H_0 , based on evidence from data?	$I(\hat{T} \neq T)$ <i>Type I or II Error:</i> Do we falsely reject or falsely accept H_0 ?	<i>Error Probability:</i> The test's rates of false rejection and false acceptance when applied to control problems.	Hwang et al. 1992, Berger et al. 1994, Berger 2003

* $I(\text{statement})$ denotes the indicator function: it equals 1 if the statement in parentheses is true and 0 otherwise.

Multi-resolution Replications

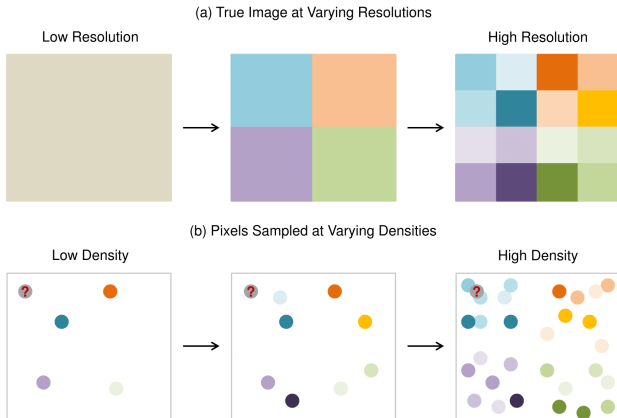
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The Problem Gets Easier But My Intervals Get Longer ?!

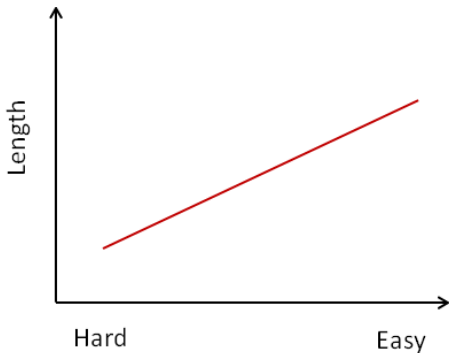
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Summary





When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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Summary

Precision as Function of Multiple Features (Basu 1964)

(X_i, Y_i) bivariate standard normal with unknown correlation θ

When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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Summary

Precision as Function of Multiple Features (Basu 1964)

(X_i, Y_i) bivariate standard normal with unknown correlation θ

- Fact 1: X_i, Y_i marginally ancillary, not jointly ancillary.

When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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Precision as Function of Multiple Features (Basu 1964)

(X_i, Y_i) bivariate standard normal with unknown correlation θ

- Fact 1: X_i, Y_i marginally ancillary, not jointly ancillary.
- Fact 2: As $\|X\|$ or $\|Y\|$ increases, precision for θ increases.

When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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Option 1: Evaluate uncertainty of $\hat{\theta}$ (MLE) *unconditionally*.
Construct pivot (using inverse CDF) and invert into CI.

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- Achieves **exact, unconditional** coverage.



When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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Precision as Function of Multiple Features (Basu 1964)

(X_i, Y_i) bivariate standard normal with unknown correlation θ

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Option 1: Evaluate uncertainty of $\hat{\theta}$ (MLE) *unconditionally*.
Construct pivot (using inverse CDF) and invert into CI.

- Achieves **exact, unconditional** coverage.

Option 2: Evaluate uncertainty of $\hat{\theta}$ *conditional* on $\|X\|$.

- But what about the effect of $\|Y\|$ on precision?

A Heterogeneous Population of Datasets

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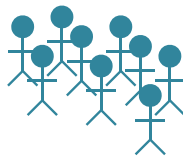
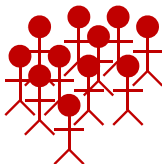
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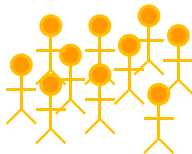
Low $\|X\|$

High $\|X\|$

Low $\|Y\|$



High $\|Y\|$



Here's Where Resolution Helps Us Reason...

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Summary

A Regression Perspective

- As $\|X\|$ increases, precision of $\hat{\theta}$ increases.

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A Regression Perspective

- As $\|X\|$ increases, precision of $\hat{\theta}$ increases.
- As $\|Y\|$ increases, precision of $\hat{\theta}$ increases.

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A Regression Perspective

- As $\|X\|$ increases, precision of $\hat{\theta}$ increases.
- As $\|Y\|$ increases, precision of $\hat{\theta}$ increases.
- The **first order** effects of $\|X\|$ and $\|Y\|$ on precision are robust to assumptions about θ .

Here's Where Resolution Helps Us Reason...

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Basu Ex

Summary

A Regression Perspective

- As $||X||$ increases, precision of $\hat{\theta}$ increases.
- As $||Y||$ increases, precision of $\hat{\theta}$ increases.
- The **first order** effects of $||X||$ and $||Y||$ on precision are robust to assumptions about θ .

But when we condition on $||X||$ and $||Y||$...

- We also model **second order** effect: how $||X||$ and $||Y||$ *together* affect data precision (their interaction).

Here's Where Resolution Helps Us Reason...

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14/21

Xiao-Li Meng

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Basu Ex

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- **Second order** effect (interaction term) is not robust to prior assumptions about θ .
- How to account for first order effects while ignoring second order effects and do so in a *principled* way?



Fiducial's Pivotal Idea (Fraser 68, Hannig 09)

God's U Always Exists

Represent data as $X = g(\theta; U)$ where $U \sim p(U)$ is known.

$$\text{Normal :} \quad X = \theta + U \quad U \sim N(0, 1)$$

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2. Convert inference for U into inference for θ by inverting $X = g(\theta; U)$ to obtain $\theta = h(U; X)$:

$$\text{E.g. : } \theta = X - U \sim N(X, 1).$$

Fiducial Inference for Bivariate Normal

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Summary

(X_i, Y_i) bivariate normal with mean 0, var 1 and correlation θ .

- Reduce to sufficient statistics: $S_1 = \sum_i (X_i + Y_i)^2$ and $S_2 = \sum_i (X_i - Y_i)^2$.

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$$\sqrt{\frac{S_1}{4Q_1}} + \sqrt{\frac{S_2}{4Q_2}} - 2 = 0$$

and $Q_i \geq S_i/16$ for $i = 1, 2$.

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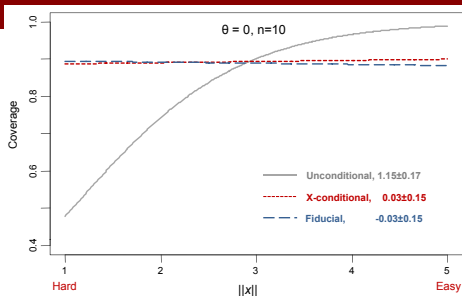
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and $Q_i \geq S_i/16$ for $i = 1, 2$.

- Inference for θ : Given Q_1, Q_2 , let $\theta = \sqrt{\frac{S_1}{4Q_1}} - 1$.

Checking Coverage and Length Conditioning on $||x||$



Checking Coverage and Length Conditioning on $||x||$

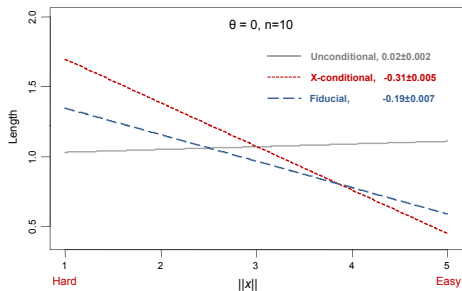
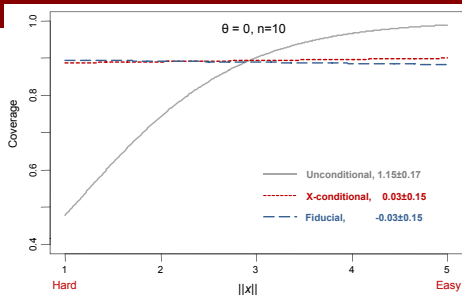
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Checking Coverage and Length Conditioning on $\|y\|$

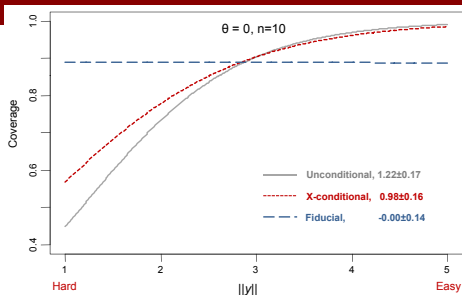
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Checking Coverage and Length Conditioning on $\|y\|$

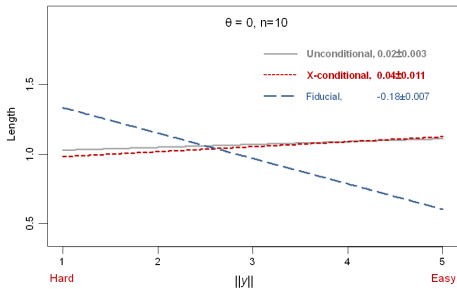
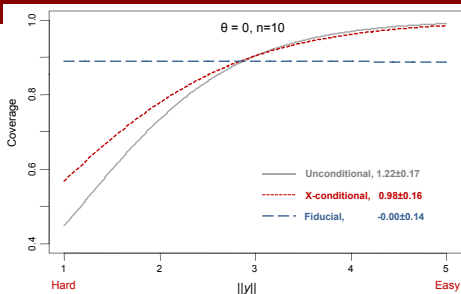
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A Fundamental Principle of Statistical Inference: Bias-Variance or Relevant-Robust Trade-off

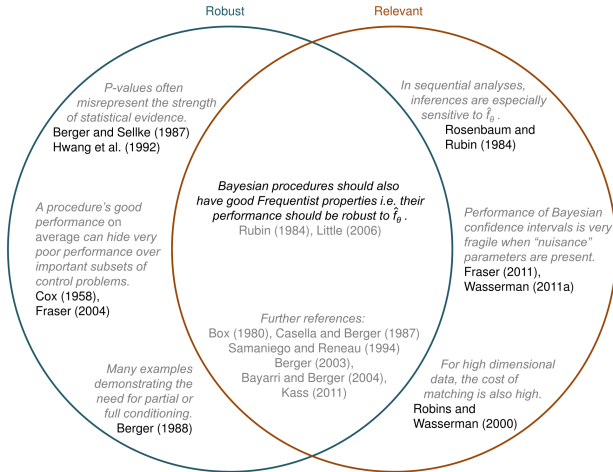
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A Unified Picture of BFF (and Inference)?

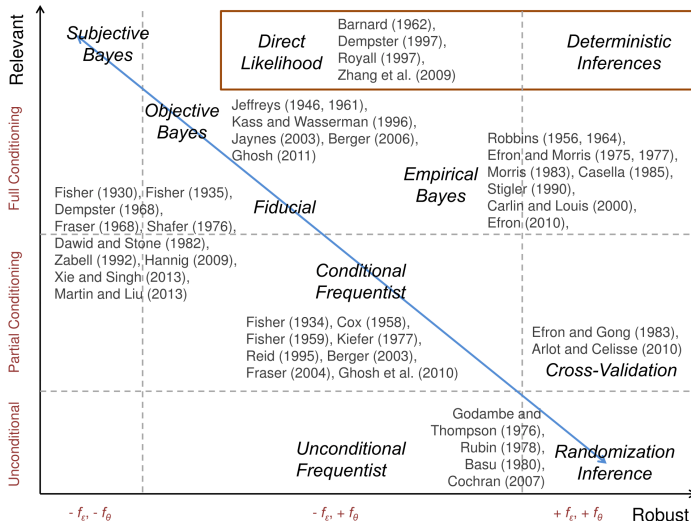
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Let's be BFF, not merely FWB ...

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