

# Probing the composite light scalar of the sextet model for dilaton fingerprints

Chik Him (Ricky) Wong

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- Sextet model:  $SU(3)$   $N_f = 2$  fermions in two-index symmetric representation
- Near conformal window  $\Rightarrow$  candidate of walking theory
- Analysis of  $M_\pi$  and  $F_\pi$  in the chiral limit using  $\chi$ PT has become difficult as  $m$  decreases in the dataset
- Previous attempts have limited success but still difficult (e.g. NLO chiral log fits, Rooted staggered  $\chi$  PT)
- Possible cause: Near conformal window, emergent  $0^{++}$  scalar can be as light as NGBs  $\Rightarrow$   $\chi$ PT needs modification
- Dilaton hypothesis: The scalar acts as a dilaton from scale symmetry breaking ( details in Julius Kuti's talk)
- Tree-level predictions of Dilaton Effective Lagrangian
  - Scaling formulae:  $M_\pi^2 F_\pi^{2-y} = C m$ , where  $C = 2B_\pi f_\pi^{2-y}$ ,  $y = 3 - \gamma$ ,  $f_\pi = F_\pi(m \rightarrow 0)$
  - Asymptotic dilaton potential  $V(\chi) \sim \chi^p$  :  $M_\pi^2 F_\pi^{2-p} = B$

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- A recent work (Appelquist et al, 2018) attempted to do these tests on dataset extracted from our published plots and claimed success
- The study was limited
  - The dataset used was outdated. New data are available and the estimates of some data entries are improved.
  - Error estimates were ball-park estimates obtained only graphically instead of using the actual data
  - Only  $\beta = 3.20$  was analyzed. Cutoff effects was not studied.
- $\Rightarrow$  The analysis is not conclusive. We need a more comprehensive study.

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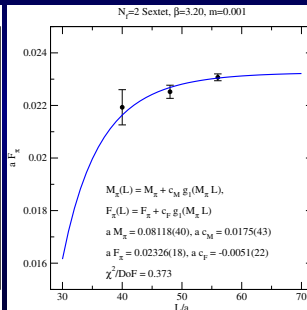
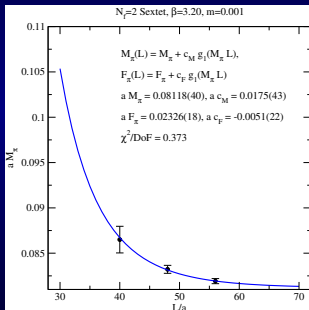
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- Finite Size Scaling
- $M_\pi$  and  $F_\pi$  values are defined in infinite volume limit, while simulations are done in finite volumes  $\Rightarrow$  Extrapolation into infinite volume limit is required
- Ansatz:

$$M_\pi(L) = M_\pi(L \rightarrow \infty) + c_M g_1(M_\pi L, \eta = N_t/N_L), \quad F_\pi(L) = F_\pi(L \rightarrow \infty) + c_F g_1(M_\pi L, \eta = N_t/N_L)$$



- In the presence of light dilaton  $M_d \sim M_\pi$ , the particle being exchanged can also be dilaton. However, the combined correction term would still be  $g_1$
- Note:  $Q = 0$  effect is ignored

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- Dilaton hypothesis

- Scaling formulae (primary test):  $M_\pi^2 F_\pi^{2-y} = C m,$
- Asymptotic dilaton potential  $V(\chi) \sim \chi^p$  :  $M_\pi^2 F_\pi^{2-p} = B$
- $M_\pi^2 F_\pi^{2-y}$  and  $M_\pi^2 F_\pi^{2-p}$  are nonlinear functions of  $(M_\pi, F_\pi, y)$  and  $(M_\pi, F_\pi, p)$  respectively
- The distribution of the error of a function  $f(M_\pi, F_\pi)$  can be approximated by a normal distribution

$$f(M_\pi, F_\pi) - f(\langle M_\pi \rangle, \langle F_\pi \rangle) \sim N(0, \nabla f^\dagger \Sigma \nabla f),$$

in which  $\Sigma$  is the covariance matrix between  $M_\pi$  and  $F_\pi$  from the FSS fit

- $\partial_{F_\pi}(M_\pi^2 F_\pi^{2-y}) = (2-y)F_\pi^{1-y} M_\pi^2$ , a function of  $y$  (same case for  $p$ )
- Problem: How to properly propagate errors in  $M_\pi$  and  $F_\pi$  into these quantities, in the presence of  $p$  and  $y$ ?

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- Generalized Maximum Likelihood Estimates
- Consider testing a hypothesis  $Y = f_{\theta}(X)$  parameterized by  $\theta$  with observable  $Y$  at  $X$ :
  - Each  $Y_i$  is measured at  $X_i$ 's with Gaussian uncertainties described by  $N(0, \sigma_{Y,i}^2)$
  - Measurements at each  $X_i$  are assumed to be independent
- The likelihood  $L(\theta; Y) \propto \exp(-\sum_i \frac{(Y_i - f_{\theta}(X_i))^2}{2\sigma_{Y,i}^2})$  can be maximized by minimizing

$$\chi^2 = \sum_i \left( \frac{Y_i - f_{\theta}(X_i)}{\sigma_{Y,i}} \right)^2,$$

- If there are more than one parameters, e.g.  $\theta = (\theta_0, \theta_1)$ , such minimization can be done alternatively holding one of them fixed until convergence.
- $\Rightarrow \sigma_{Y,i}^2$  can be functions of  $\theta$  as long as the uncertainty of  $Y_i$  can be described by  $N(0, \sigma_{Y,i}^2(\theta))$  at all values of  $\theta$

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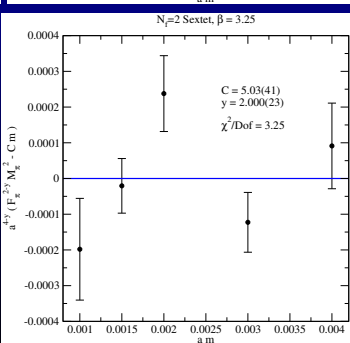
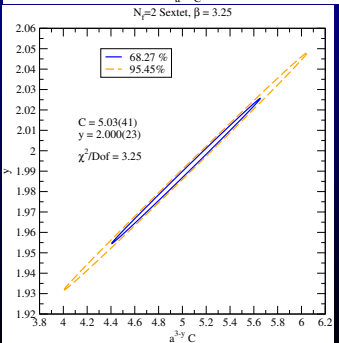
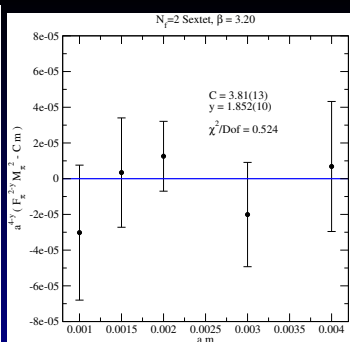
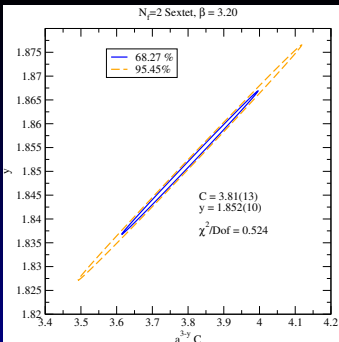
- Assuming the uncertainty of  $M_\pi^2 F_\pi^{2-y}$  is described by  $N(0, \nabla(M_\pi^2 F_\pi^{2-y})^\dagger \Sigma \nabla(M_\pi^2 F_\pi^{2-y}))$  at all values of  $y$ , (and similar for  $M_\pi^2 F_\pi^{2-p}$ ), one obtains the generalized ML estimates for  $(C, y)$  and  $(B, p)$  by minimizing  $\chi^2$  as in ordinary fitting procedures



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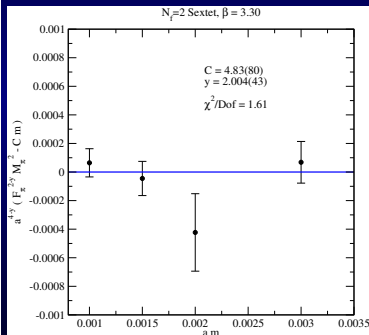
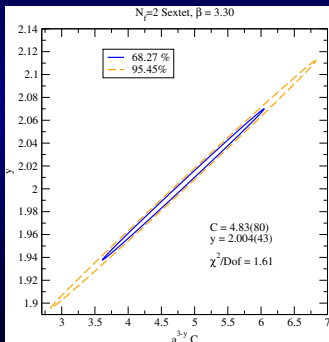
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- At  $\beta = 3.30$ , we do not have enough data for infinite volume extrapolation, but we still attempted to fit using the largest volumes available. ( $64^3 \times 96$  for  $m = 0.001 - 0.002$  and  $48^3 \times 96$  for  $m = 0.003$ )

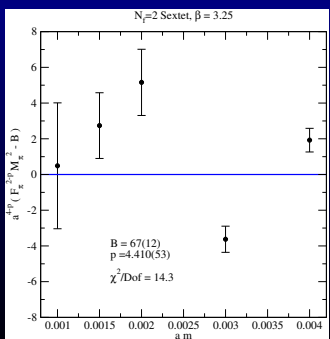
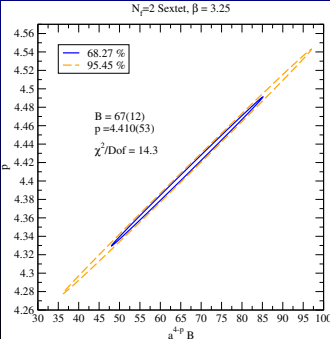
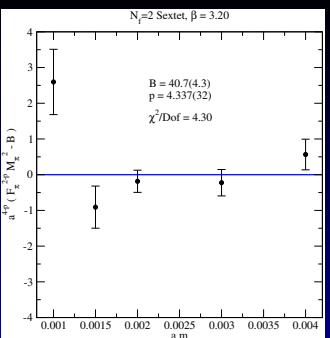
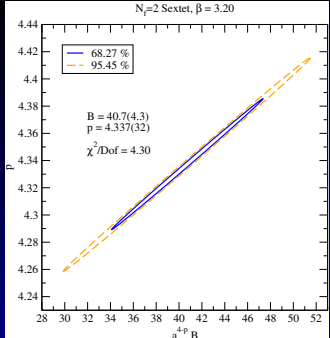


- The  $(C, y)$  fits seem to work with acceptable  $\chi^2/Dof$  at all  $\beta$ 's

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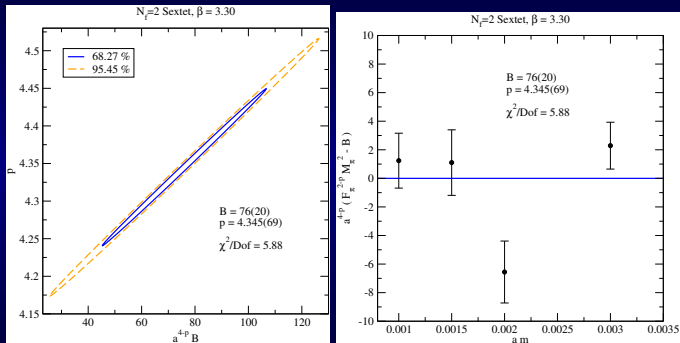
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- Again, we fitted  $\beta = 3.30$  without infinite volume extrapolation



- The  $(B, p)$  fits fail, especially at  $\beta = 3.25$

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- Our data are consistent with the scaling formulae ( $(C,y)$  fits) , but not with the asymptotic dilaton potential ( $(B,p)$  fits)
- Possible explanations:
  - (a) The ansatz is problematic. e.g. The asymptotic dilaton potential may not be of the form  $\chi^p$ .
  - (b) Some Normal distribution approximations in our analysis do not hold
- Would the  $(B,p)$  fits work if some of the Normal distribution approximations in our analysis are improved?  $\Rightarrow$  Monte Carlo and Markove Chain Monte Carlo can help

# Monte Carlo Analysis

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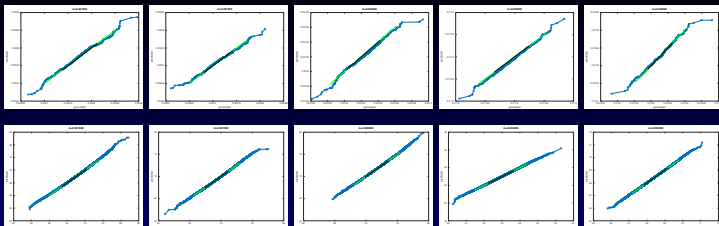
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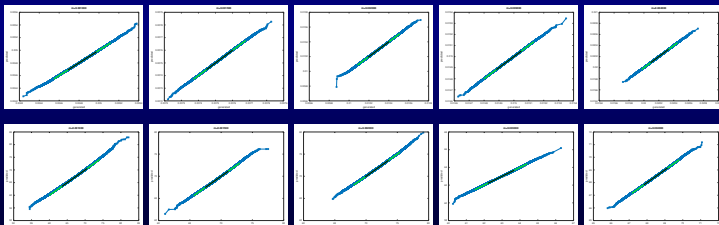
Conclusion

- Assumption: the uncertainty of  $M_\pi^2 F_\pi^{2-y}$  is described by  $N(0, \nabla(M_\pi^2 F_\pi^{2-y})^\dagger \Sigma \nabla(M_\pi^2 F_\pi^{2-y}))|_m$  at all values of  $y$  at each  $m$  (and similar for  $M_\pi^2 F_\pi^{2-p}$ )
- In order to improve beyond Gaussian by expansion, one needs higher moments than  $\Sigma_m$
- With only  $\Sigma_m$ , one can still do the following :
  - Approximate the distribution of  $\theta_m = \{M_\pi, F_\pi, c_M, c_F\}|_m$  as a multivariant Gaussian distribution  $N(\langle\theta\rangle_m, \Sigma_m)$  at each  $m \Rightarrow$  Numerically generate such distribution
  - Compute  $M_\pi^2 F_\pi^{2-y}$  for the range of  $y$ 's suggested by the fitted values (similar for  $M_\pi^2 F_\pi^{2-p}$ )
  - Compare the distributions with the ones predicted by  $N(\langle M_\pi \rangle^2 \langle F_\pi \rangle^{2-y}, \nabla(M_\pi^2 F_\pi^{2-y})^\dagger \Sigma \nabla(M_\pi^2 F_\pi^{2-y}))|_m$  (similar for  $M_\pi^2 F_\pi^{2-p}$ )
  - A simple way is to look at Quantile-Quantile plots. The closer to the  $y = x$  line the more similar these two distributions are.

$\beta = 3.20$ : First row:  $M_\pi^2 F_\pi^{2-y_{\text{fit}}}$  Second row:  $M_\pi^2 F_\pi^{2-p_{\text{fit}}}$



$\beta = 3.25$ : First row:  $M_\pi^2 F_\pi^{2-y_{\text{fit}}}$  Second row:  $M_\pi^2 F_\pi^{2-p_{\text{fit}}}$



- They are almost identical and normal up to 2 sigma range.
- The uncertainties of fitted  $p$  and  $y$  are narrow, so only the fitted values are used here for illustration. Yet this level of resemblance holds at broader range of values.

# Markov Chain Monte Carlo (MCMC) Analysis

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- Assumption of Monte Carlo analysis: At each  $m$ ,  
 $\theta_m = (M_\pi(L \rightarrow \infty), F_\pi(L \rightarrow \infty), c_M, c_F)|_m$  distributes as Gaussian  
with covariance matrix  $\Sigma_m$

- Actual assumption in Maximum Likelihood estimate:

$$\begin{aligned} & (M_\pi(L_i) - [M_\pi(L \rightarrow \infty) + c_M g_1(M_\pi L, \eta)])|_m \\ & (F_\pi(L_i) - [F_\pi(L \rightarrow \infty) + c_F g_1(M_\pi L, \eta)])|_m \end{aligned}$$

follows  $N(0, \Sigma_m)$  at each  $m$

- In Frequentist picture, it is the likelihood  $L(\theta_m; \{M_\pi(L_i), F_\pi(L_i)\}_m)$   
In Bayesian picture, this distribution is equivalent to the posterior  
distribution with uniform prior  $P(\theta) \propto 1$

$$P(\theta_m | \{M_\pi(L_i), F_\pi(L_i)\}_m) \propto e^{-\chi_m^2/2} P(\theta) \propto e^{-\chi_m^2/2}$$

- This distribution can be generated by Markov Chain Monte Carlo (MCMC) method



# Markov Chain Monte Carlo (MCMC)

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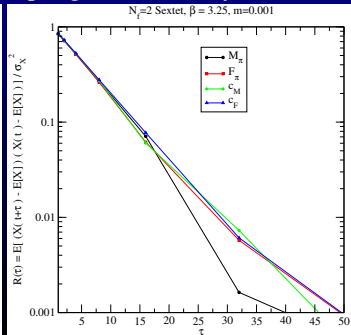
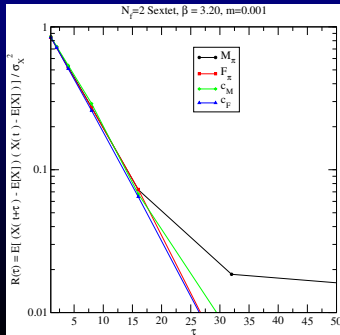
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- Markov Chain Monte Carlo algorithms sample a desired distribution by constructing a Markov Chain with such distribution as the equilibrium distribution
- A simple implementation is the familiar Metropolis-Hastings algorithm using a Gaussian proposal density
- Once we obtain the distribution, we can check again with  $N ( \langle M_\pi \rangle^2 \langle F_\pi \rangle^{2-y}, \nabla ( M_\pi^2 F_\pi^{2-y} )^\dagger \Sigma \nabla ( M_\pi^2 F_\pi^{2-y} ) ) |_m$  (similar for  $M_\pi^2 F_\pi^{2-p}$ )
- Autocorrelation is checked and sampling is taken every 64<sup>th</sup>



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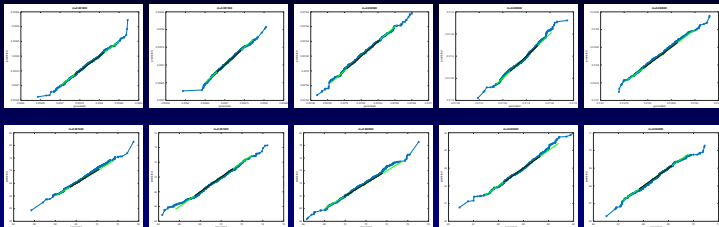
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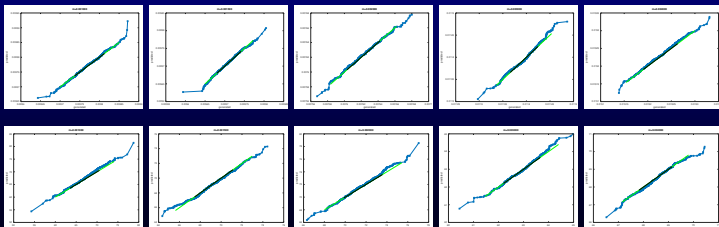
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$\beta = 3.20$ : First row:  $M_\pi^2 F_\pi^{2-y_{\text{fit}}}$  Second row:  $M_\pi^2 F_\pi^{2-p_{\text{fit}}}$

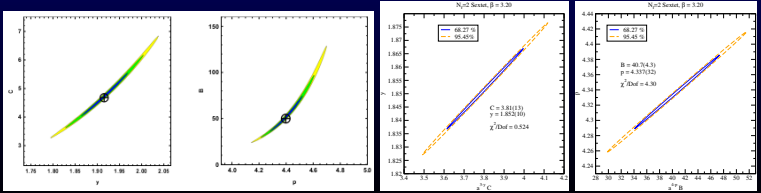


$\beta = 3.25$ : First row:  $M_\pi^2 F_\pi^{2-y_{\text{fit}}}$  Second row:  $M_\pi^2 F_\pi^{2-p_{\text{fit}}}$



- The distribution is normal up to 2 sigmas again

- All three reproduces one another's results  $\Rightarrow$  Our analysis procedures are robust and correct. However,
- Appelquist et al, 2018: Analysis of extracted data  $\beta = 3.20$ 
  - $y = 1.9(1)$ ,  $C = 4.7$ ,  $\chi^2/\text{DoF} = 0.19$
  - $p = 4.4(3)$ ,  $B = 51$ ,  $\chi^2/\text{DoF} = 0.64$
- Our analysis on our data:  $\beta = 3.20$ 
  - $y = 1.852(10)$ ,  $C = 3.81(13)$ ,  $\chi^2/\text{DoF} = 0.524$
  - $p = 4.337(32)$ ,  $B = 40.7(4.3)$ ,  $\chi^2/\text{DoF} = 4.30$
- The two results differ, especially in the  $(B,p)$  test. Possible reasons:
  - Their extracted data were outdated and inaccurate. We use our exact up-to-date dataset.
  - Some systematic errors may have been overlooked
  - Further investigation is needed
- Their work did not check  $\beta = 3.25$  data in which  $(B,p)$  test fails more in our analysis



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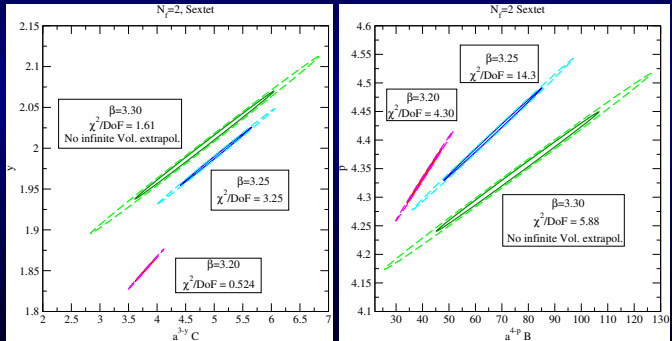
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- The dilaton hypothesis is tested in the sextet model using several different comprehensive analysis methods. Three  $\beta$  values are analyzed in order to probe cutoff effects.  $\beta = 3.30$  is incomplete due to the lack of infinite volume extrapolation.



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- The scaling formulae  $M_\pi^2 F_\pi^{2-\gamma} = C$  is consistent with our current data. However,  $\gamma = 3 - y$  should be cutoff-dependent  $\Rightarrow$  what scale does  $y$  correspond to ?
- The assumption of asymptotic dilaton potential form  $V(\chi) \sim \chi^p$  does not pass the test of  $M_\pi^2 F_\pi^{2-p} = B$  in our data
  - Possible explanations:
    - It is a tree-level prediction. The significance of 1-loop correction is unknown. In principle it is needed in order to be consistent with the FSS which probes an 1-loop correction effect.
    - Some terms may be missing in the dilaton potential
    - At small  $m$ , our data is known to be mostly frozen at  $Q = 0$ . Its effect is not taken into account in our analysis
- Cutoff effects need to be investigated

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