

# Efficient Neutrino Oscillation Parameter Inference with Gaussian Process

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New Perspective 2.0 - 2020

# Neutrino Oscillation

- Neutrinos: 2 kinds of states, each comes with 3 types
  - Flavor States ( $\nu_e, \nu_\mu, \nu_\tau$ ) — what we observed
  - Mass Eigenstates ( $\nu_1, \nu_2, \nu_3$ ) — what in between observations
- Principle of superposition connects them via PMNS matrix, i.e.

$$\begin{bmatrix} \nu_e \\ \nu_\mu \\ \nu_\tau \end{bmatrix} = U_{PMNS} \begin{bmatrix} \nu_1 \\ \nu_2 \\ \nu_3 \end{bmatrix}$$
$$U_{PMNS} = \begin{pmatrix} 1 & 0 & 0 \\ 0 & \cos \theta_{23} & \sin \theta_{23} \\ 0 & -\sin \theta_{23} & \cos \theta_{23} \end{pmatrix} \begin{pmatrix} \cos \theta_{13} & 0 & \sin \theta_{13} e^{-i\delta} \\ 0 & 1 & 0 \\ -\sin \theta_{13} e^{i\delta} & 0 & \cos \theta_{13} \end{pmatrix} \begin{pmatrix} \cos \theta_{12} & \sin \theta_{12} & 0 \\ -\sin \theta_{12} & \cos \theta_{12} & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

# Neutrino Oscillation

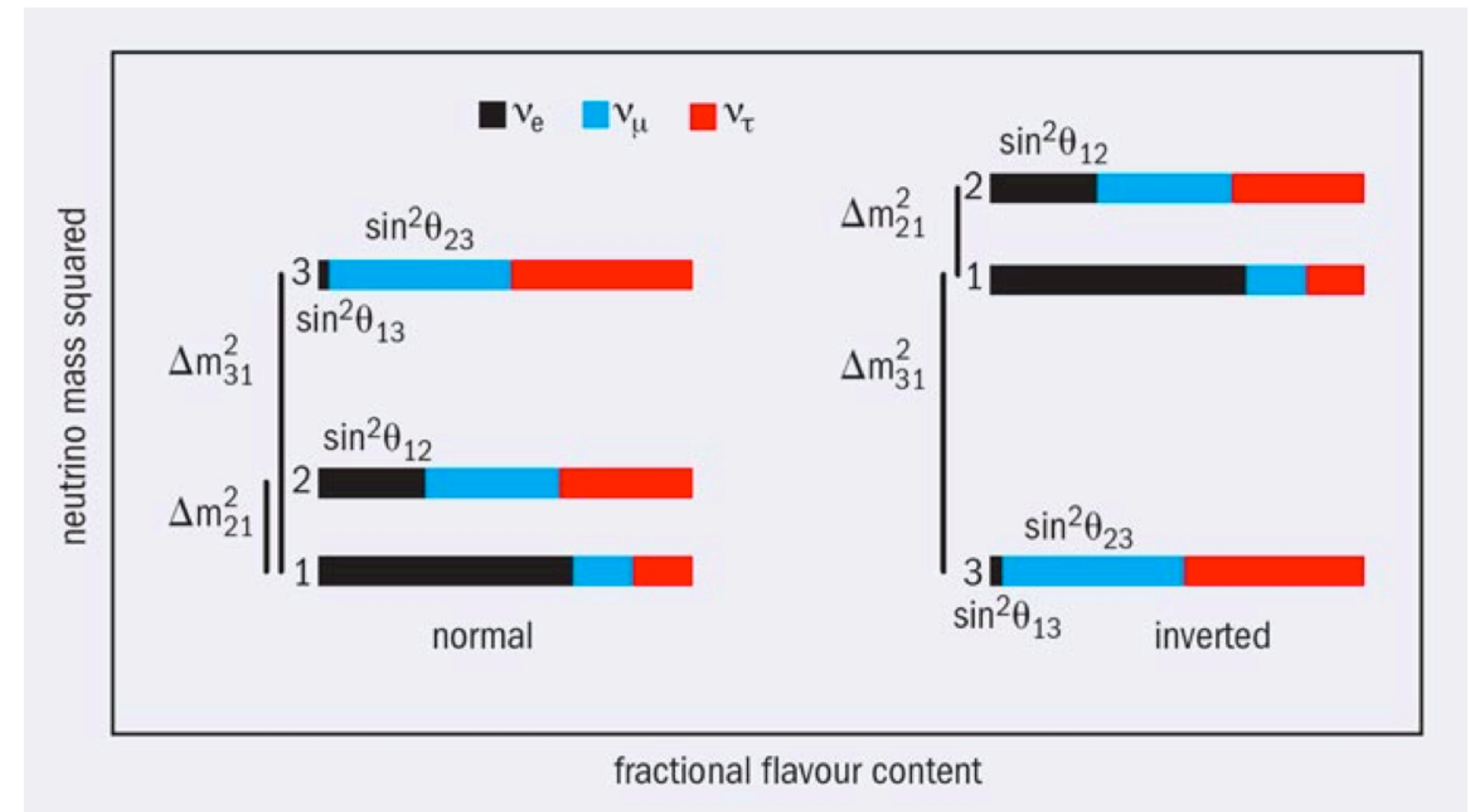
- For neutrino propagating in vacuum,

$$P(\nu_\alpha \rightarrow \nu_\beta) = \delta_{\alpha\beta} - 4 \sum_{i>j}^3 \Re(U_{\alpha i}^* U_{\beta i} U_{\alpha j} U_{\beta j}^*) \sin^2\left(\frac{\Delta m_{ij}^2 L}{4E_\nu}\right) + 2 \sum_{i>j}^3 \Im(U_{\alpha i}^* U_{\beta i} U_{\alpha j} U_{\beta j}^*) \sin\left(\frac{\Delta m_{ij}^2 L}{4E_\nu}\right)$$

- Broadly, solar experiments give handle on (21) parameters, reactor experiments for  $\theta_{13}$
- Long baseline (LBL) experiments gives handle on (32) parameters
  - $P(\nu_\mu \rightarrow \nu_\mu)$  is sensitive to  $\sin^2(2\theta_{23})$  and  $|\Delta m_{32}^2|$
  - Non-zero  $\theta_{13}$  opens up  $P(\nu_\mu \rightarrow \nu_e)$  channel, sensitive to  $\delta_{CP}$ ,  $\theta_{23}$  octant, and  $\text{sgn}(\Delta m_{32}^2)$

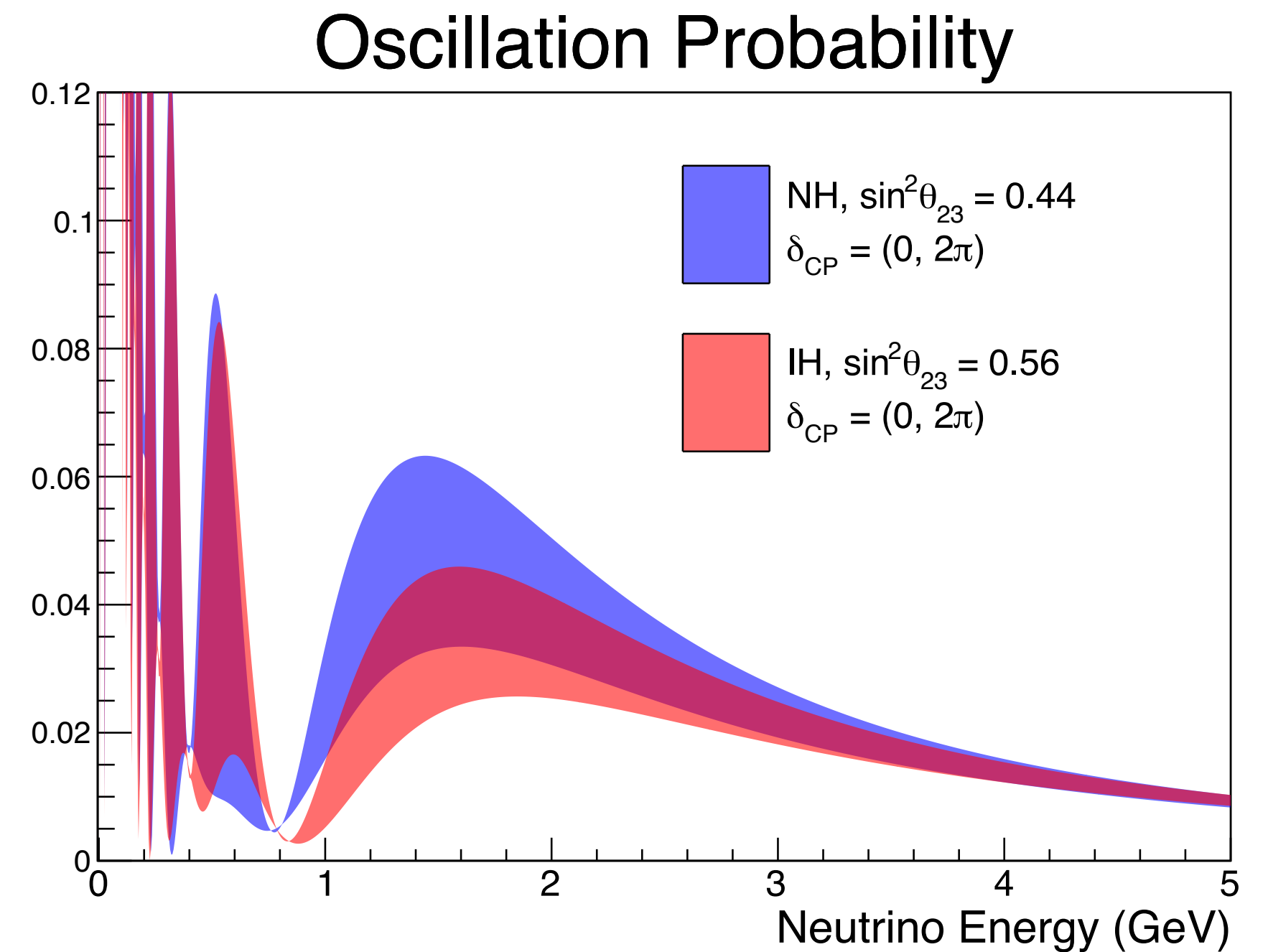
# Neutrino Oscillation

- In LBL experiments, we want to know if
  - $\Delta m_{32}^2 > 0$  or  $< 0$  ? (Normal Hierarchy or Inverted Hierarchy)
    - Has implication for neutrino mass measurements
  - Octant of  $\theta_{23}$  or  $\theta_{23} = 45^\circ$  ?
  - $\sin(\delta_{CP}) \neq 0$  ?
    - Indicate whether CP-violation exists. (matter-antimatter asymmetry)



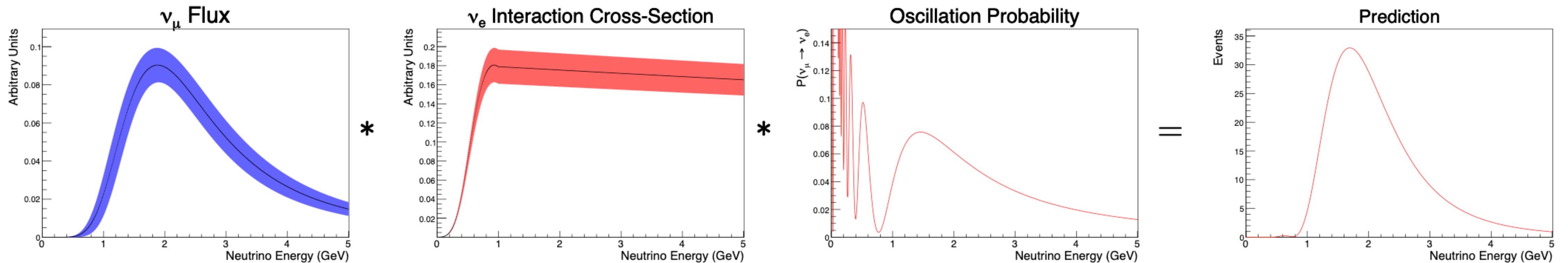
# Statistic Issues

- Oscillation parameters usually measured via the Maximum Likelihood Estimation (MLE) using the PMNS model and comparing it to the observation, such as the energy spectrum.
- However, LBL experiments (T2K, NOvA) only collect a handful of statistics over years of operation.
- Oscillation Probability have complicated dependence on multiple parameters  $\rightarrow$  difficult to delineate
- Therefore, Confidence Intervals are hard to construct



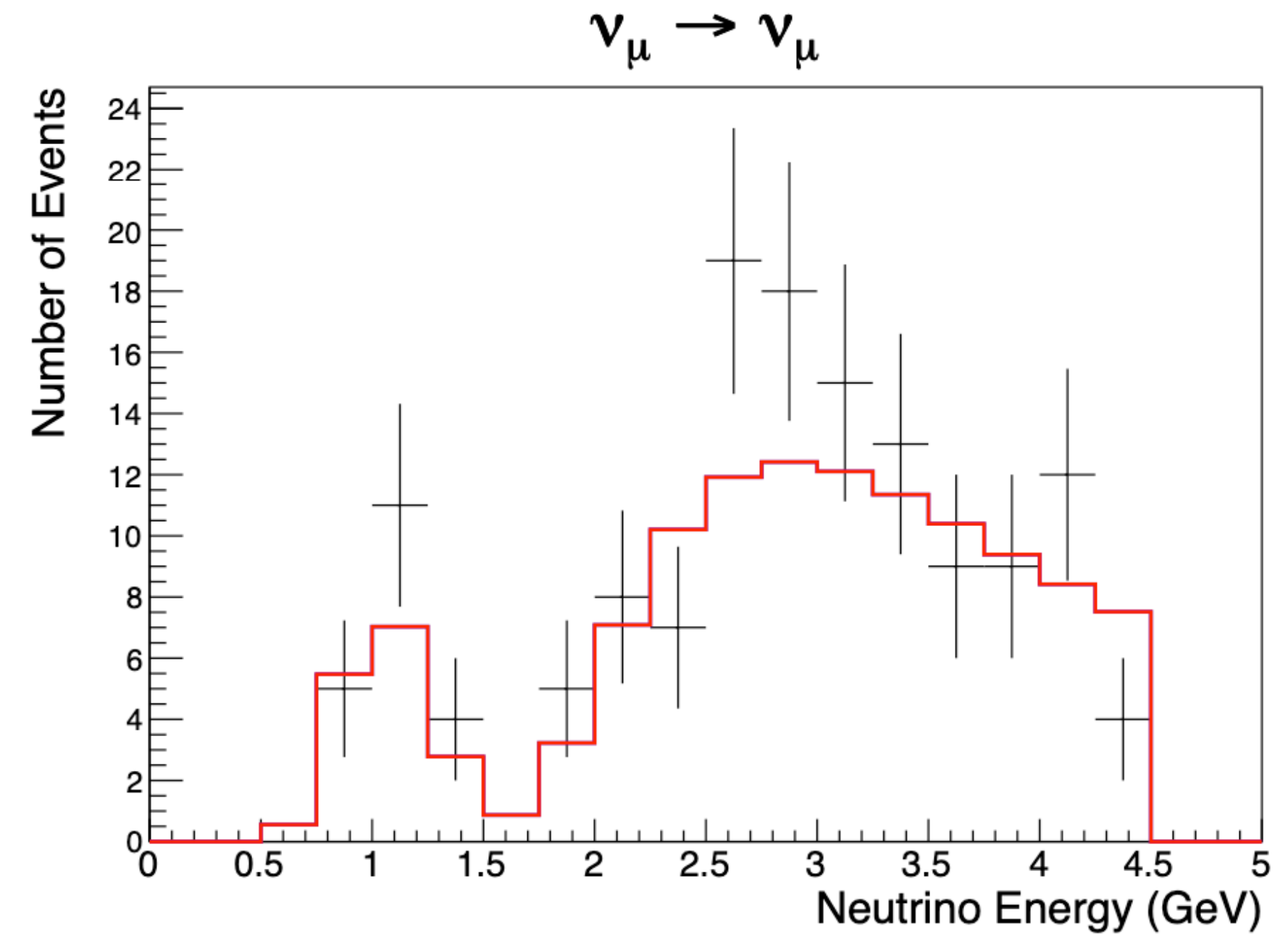
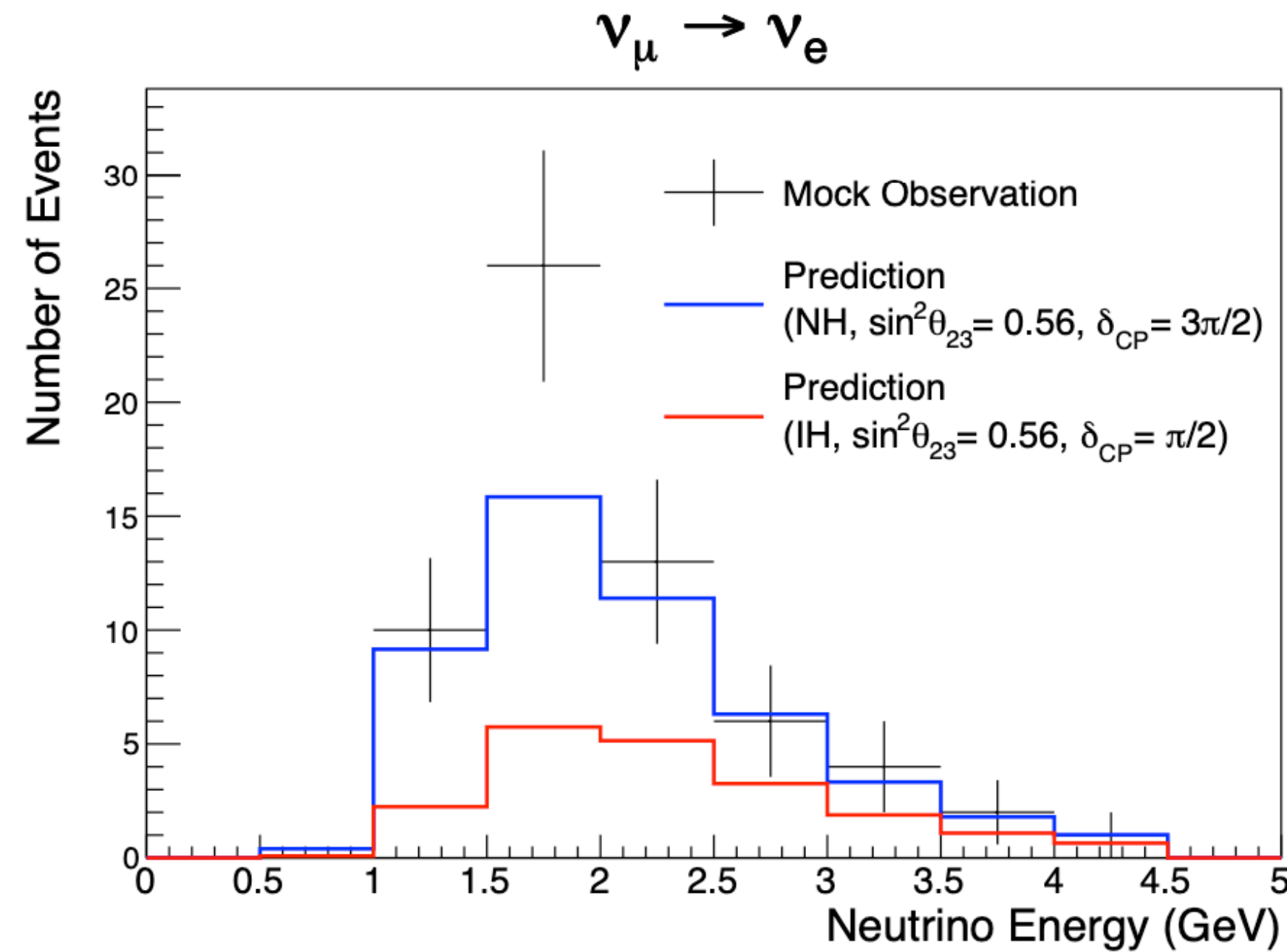
# Toy Experiment

- Modeled on NOvA. Set, baseline,  $L = 810\text{km}$  with  $\nu_\mu$  flux peaking at 2 GeV
- Add, 10% normalization error on flux and cross-section model.
- Get,  $\nu_\mu \rightarrow \nu_e$  prediction by multiplying toy shapes for flux, cross-section and oscillation probability.



- Simulate both appearance,  $\nu_\mu \rightarrow \nu_e$ , and disappearance,  $\nu_\mu \rightarrow \nu_\mu$ .

# Toy Experiment



- Toy data ( $\vec{x}$ ) from Poisson variation with some chosen oscillation parameters.
- Including both oscillation parameters,  $\theta$ , and nuisance parameters (flux and cross-sections error),  $\delta$ .
- Best fit ( $\hat{\theta}, \hat{\delta}$ ) is found by minimizing negative log-likelihood over every energy bins,  $i$

$$-2 \log L(\theta, \delta) = -2 \sum_{i \in I} \log \text{Pois}(x_i; v(\theta, \delta)_i) - \sum_{i \in I} x_i + \sum_{i \in I} v(\theta, \delta)_i + \delta^2$$



# Confidence Interval

- Typically, using Likelihood Ratio Test (LRT) to estimate confidence interval

$$\Delta\chi^2 = -2 \log \frac{L(\theta_0)}{\arg \max_{\theta} L(\theta)}$$

- In asymptotic case, test statistic:  $\Delta\chi^2 \sim \chi_k^2$  (Wilks Theorem)

**Table 38.2:** Values of  $\Delta\chi^2$  or  $2\Delta \ln L$  corresponding to a coverage probability  $1 - \alpha$  in the large data sample limit, for joint estimation of  $m$  parameters.

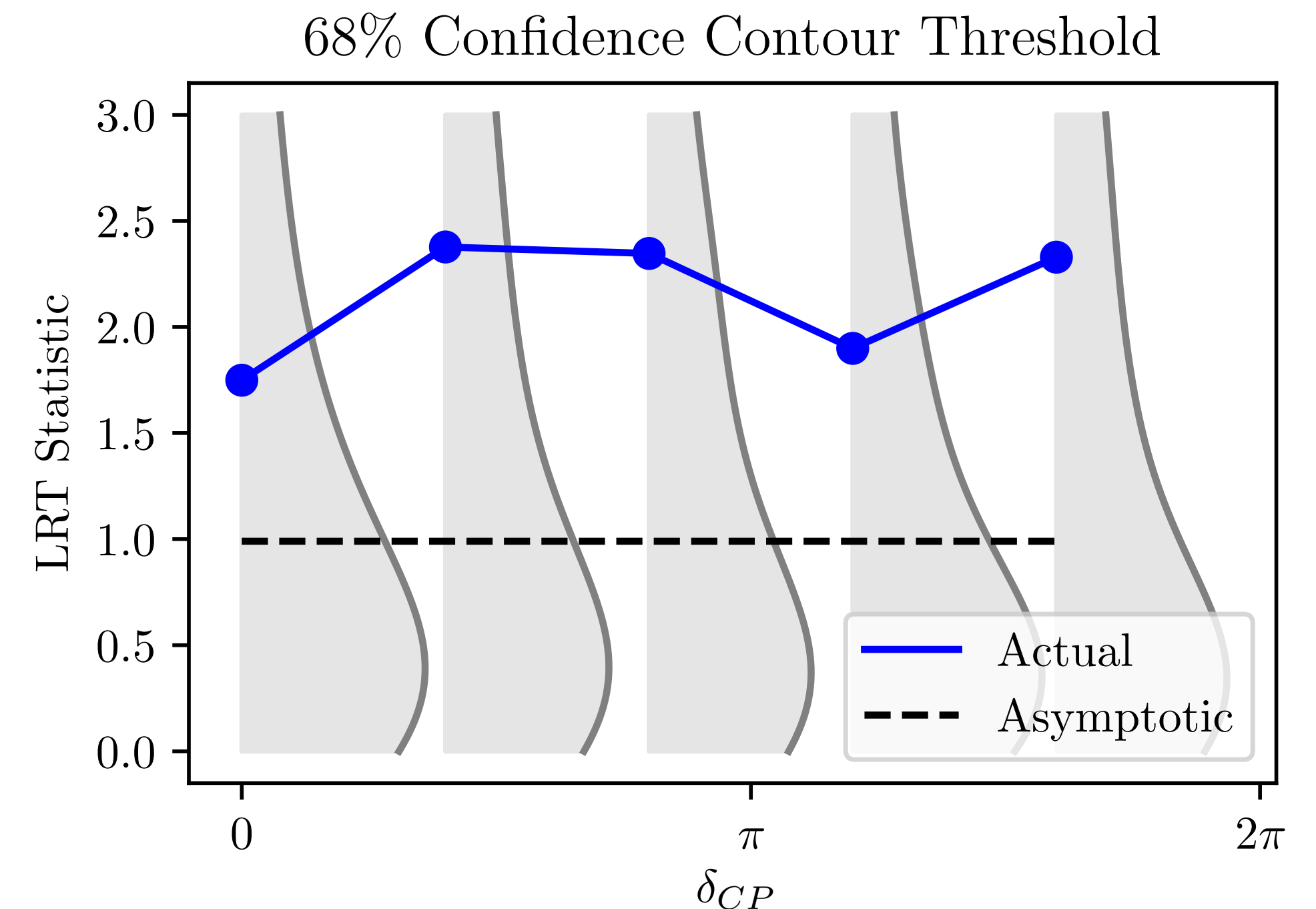
$(1 - \alpha)$ (%)	$m = 1$	$m = 2$	$m = 3$
68.27	1.00	2.30	3.53
90.	2.71	4.61	6.25
95.	3.84	5.99	7.82
95.45	4.00	6.18	8.03
99.	6.63	9.21	11.34
99.73	9.00	11.83	14.16

From the PDG Review on Statistics



# Feldman - Cousins

- Due to the small sample size in neutrino data and physical boundaries on the oscillation parameters, the asymptotic distribution is unreliable
- Explicitly simulate  $\Delta\chi^2$  distribution using lots of pseudo-experiments
- Find p-value associated with  $\Delta\chi^2_{data}$  for each point
- In practice, FC conducts a grid-search over the entire parameter space with many toy Monte-Carlo (MC) — Time Consuming
- Want a refined algorithm.
  - Approximating FC P-value surface non-parametrically using only a fraction of grid points



# Gaussian Process

- A Gaussian Process (GP) is a special case of Bayesian learning.
- Technically, GP can be specified by a mean function,  $\mu(x)$  and a covariance function (kernel),  $k(x, x')$
- Assume a collection of random models with certain probability (Priors with mean and standard deviation)

$$\begin{pmatrix} f(x) \\ f(x') \end{pmatrix} \sim \mathcal{N}\left(\begin{bmatrix} \mu(x) \\ \mu(x') \end{bmatrix}, \begin{bmatrix} k(x, x) & k(x, x') \\ k(x, x') & k(x', x') \end{bmatrix}\right)$$

- Observed data update Priors  $\rightarrow$  Posteriors (Predictions for new data)

$$f(x')|f(x) \sim \mathcal{N}\left(\frac{k(x, x')}{k(x, x)}f(x), k(x', x') - \frac{k(x, x')^2}{k(x, x)}\right)$$

- Quantifies uncertainty in model estimates (Posterior mean and standard deviation)

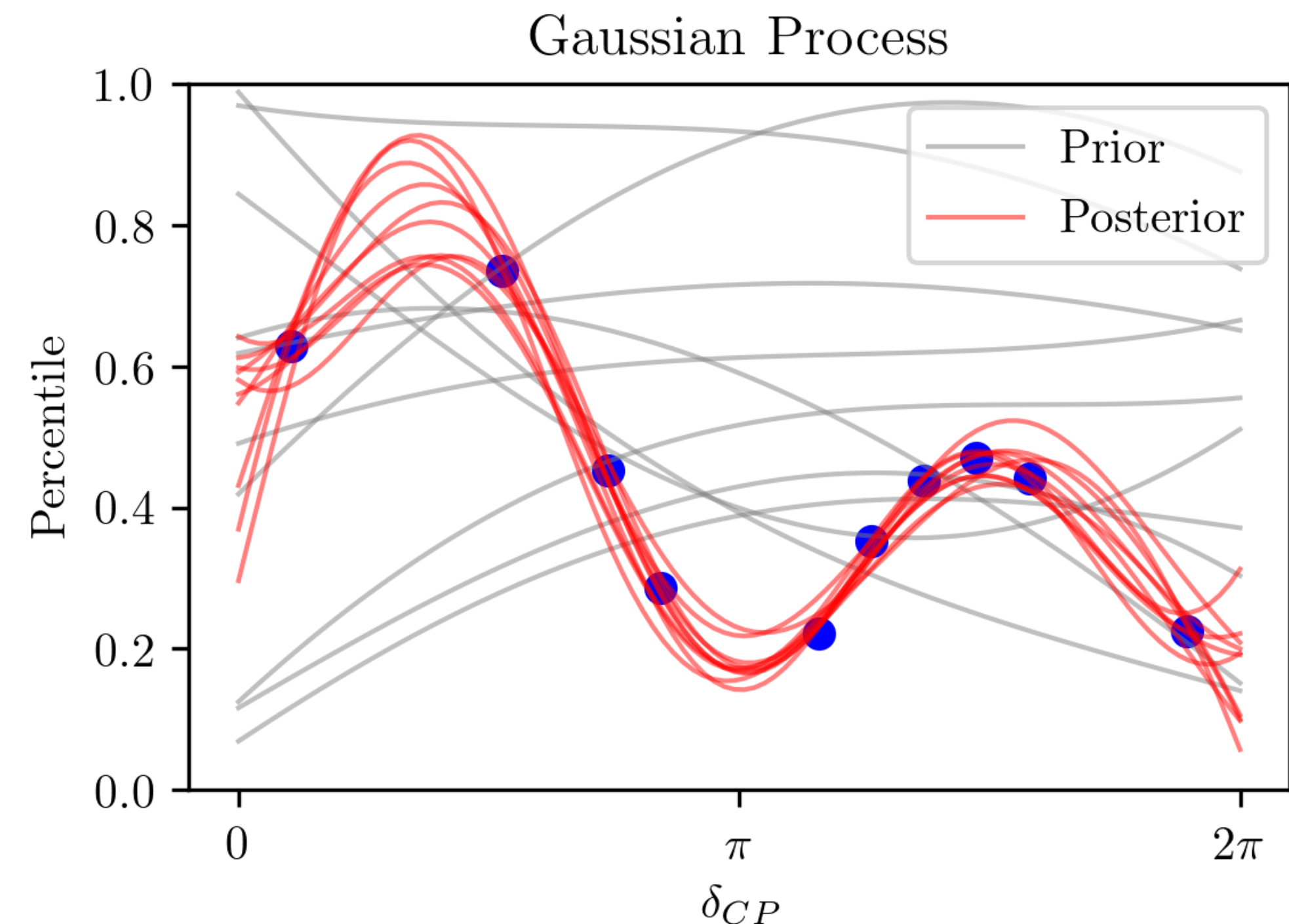
# Gaussian Process

- A common choice of GP kernel  $k$  is the squared exponential radial basis function (RBF)

$$k(x_1, x_2) = \exp\left(-\frac{(x_1 - x_2)^2}{l^2}\right)$$

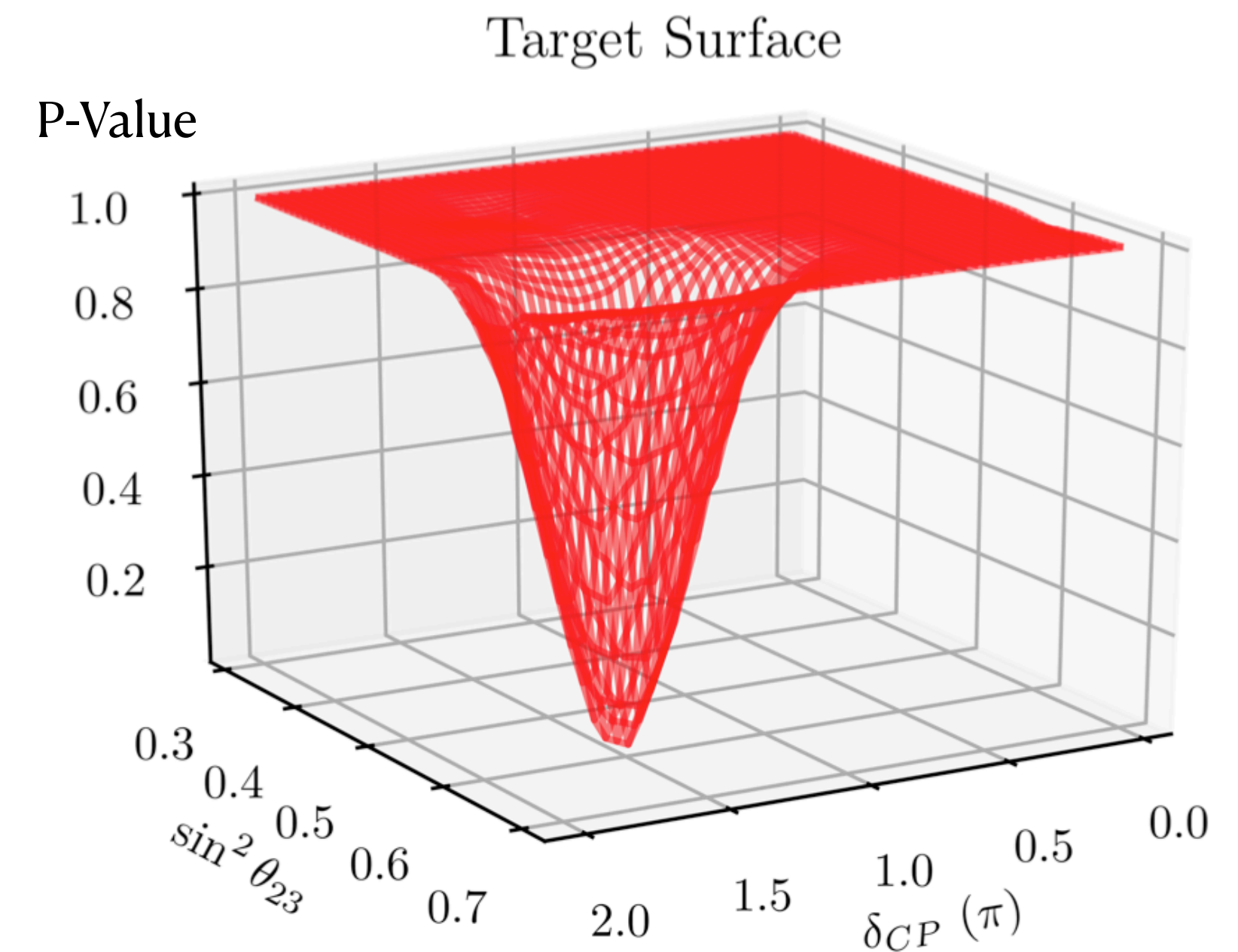
, which tells us that GP results at nearby points are highly influenced by observations at a given point while further out, they aren't.

- ***GP uses the kernel function and measured data to predict the value for an unseen point with posterior mean and posterior standard deviation***



# Gaussian Process for Feldman-Cousins Method

- Fitting a GP to target p-value surface for a given contour.
- Reduce the time-cost by throwing pseudo-experiments at some point based on approximation, instead of all points in parameter space.
- Construct Confidence Interval based on the P-value surface.



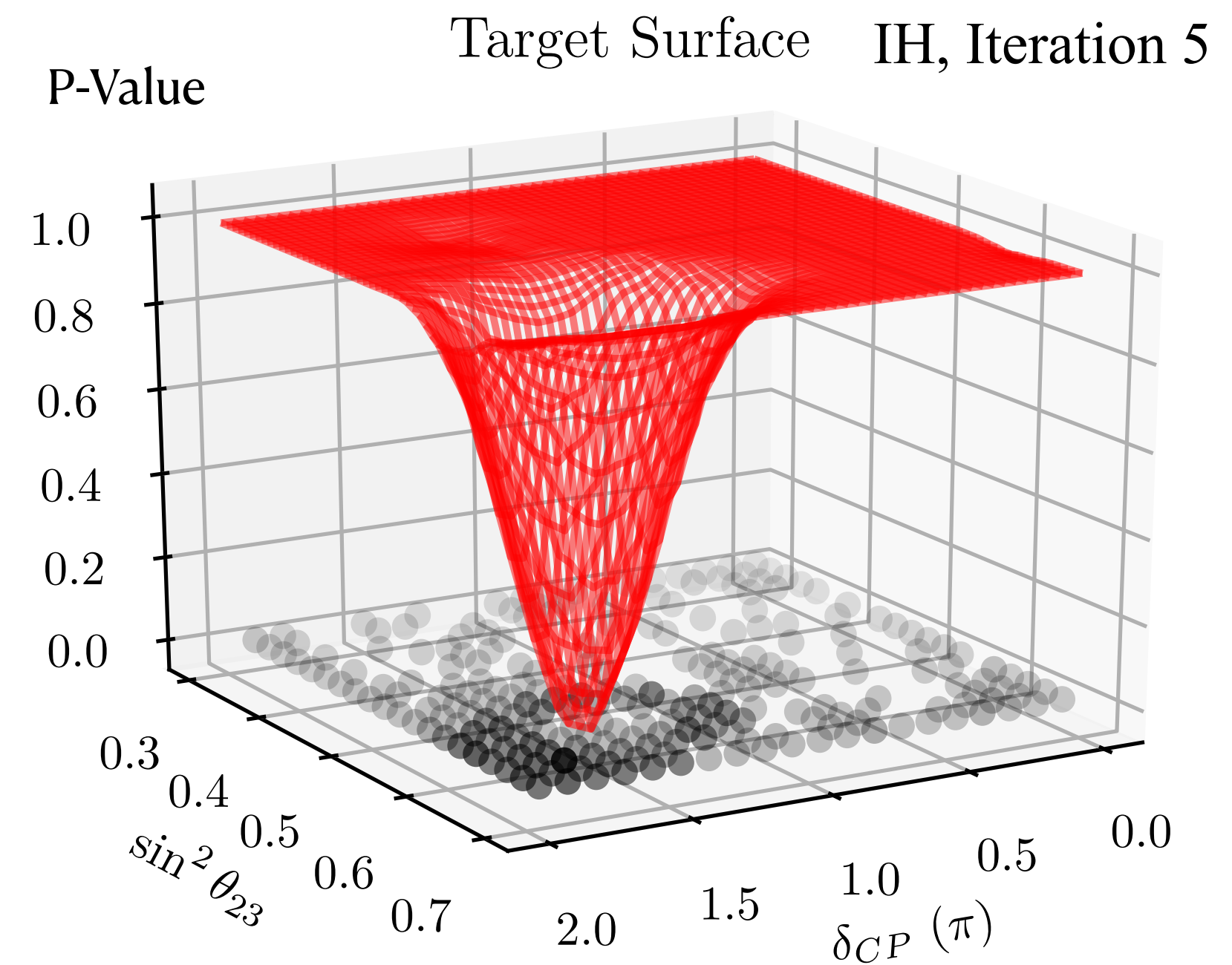
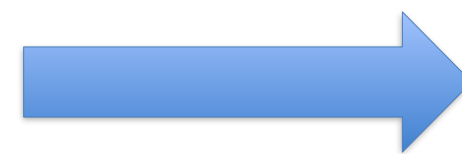
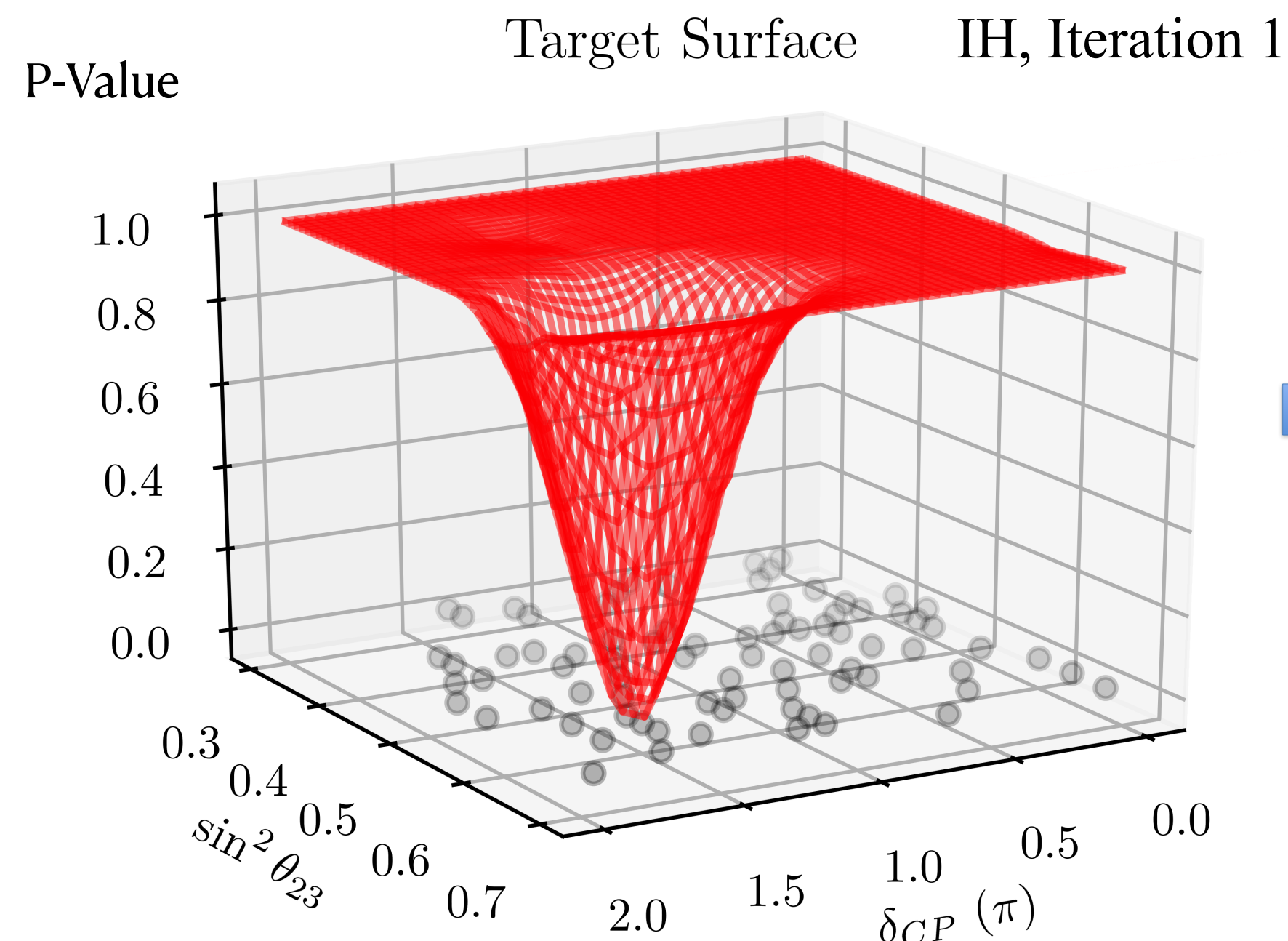


# Optimized Confidence Interval Search

- Use an acquisition function that proposed new points in  $\theta$ -space to explore based on GP approximated p-value surface

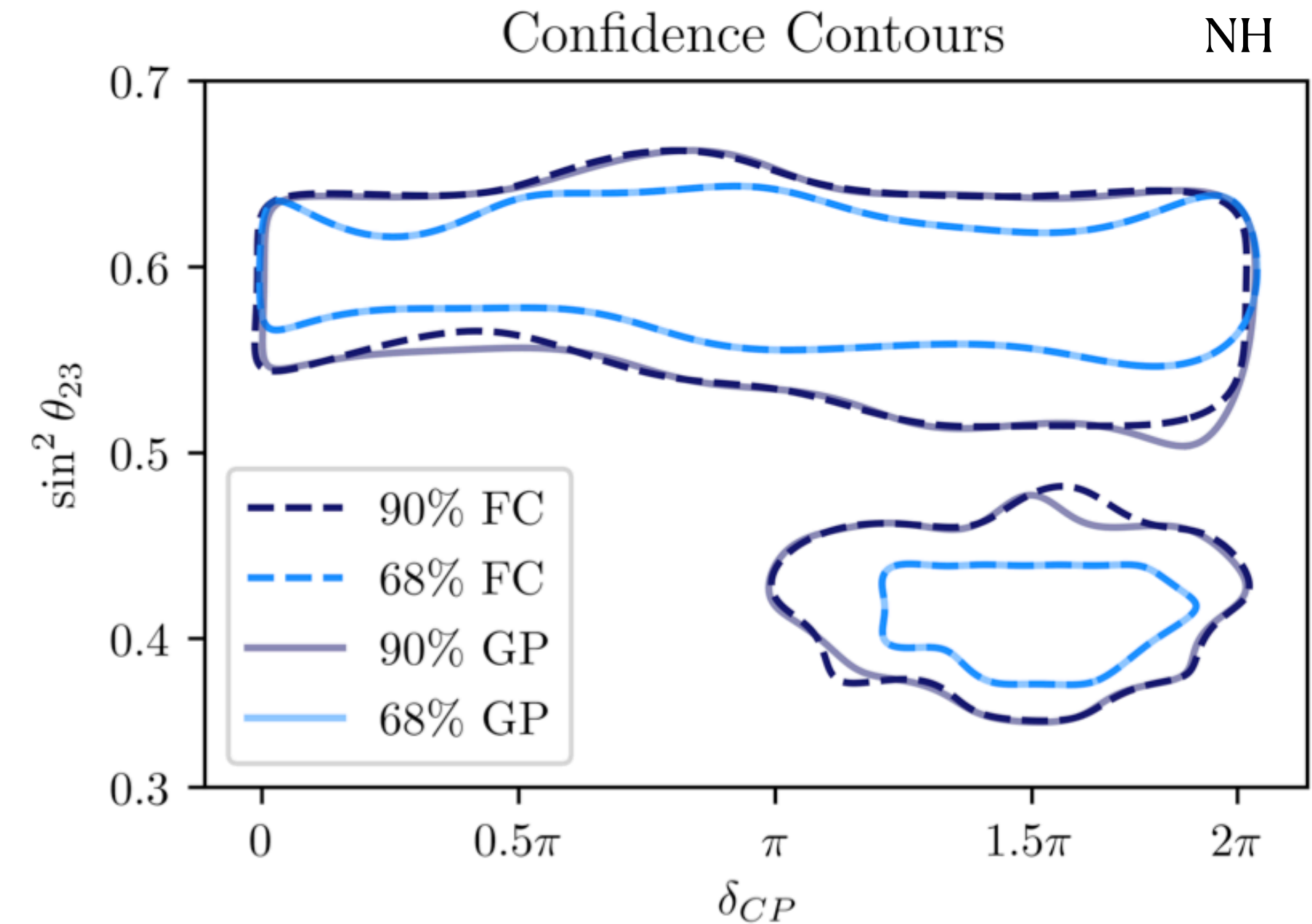
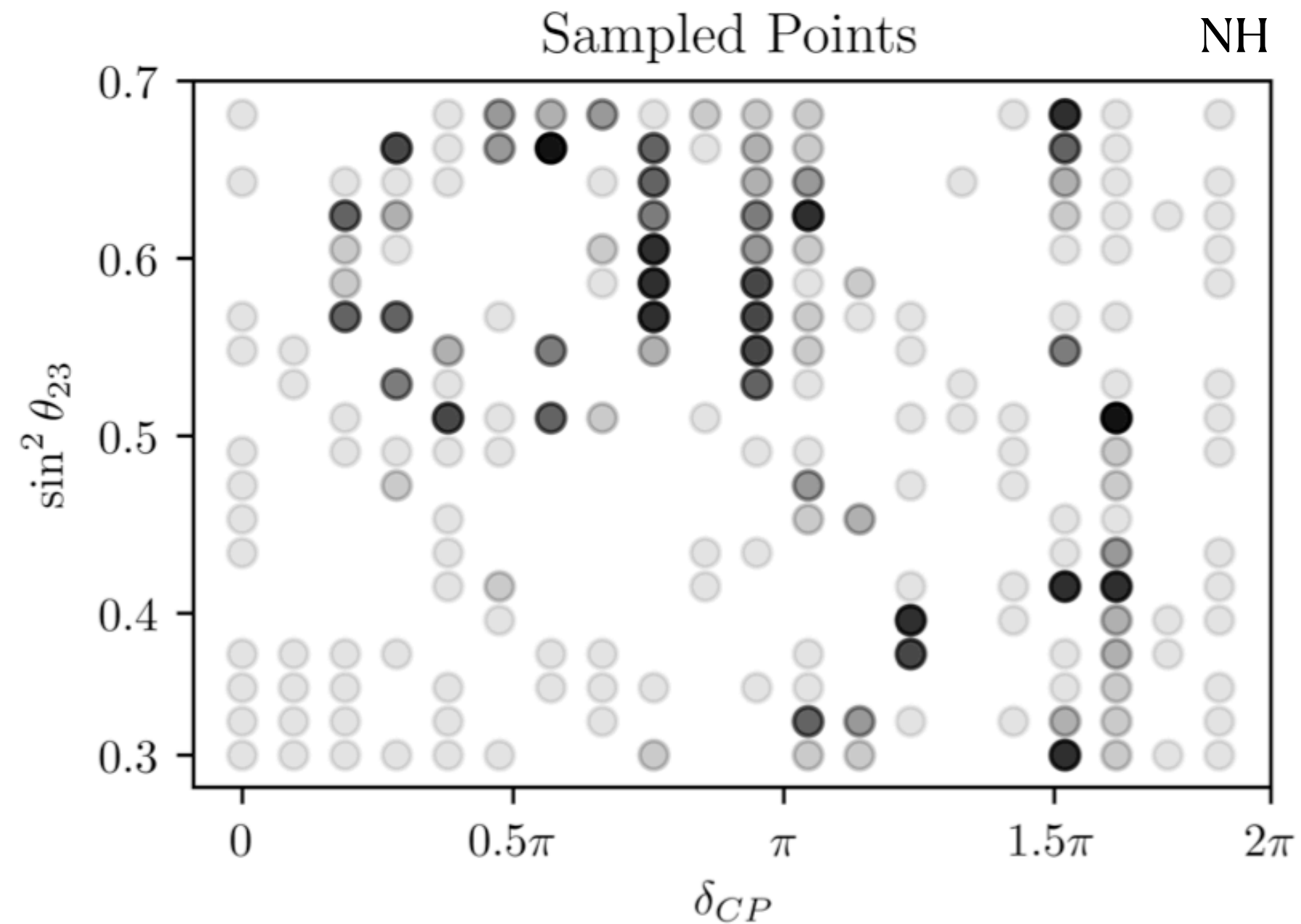
$$a(\theta) = \sum_{\alpha_i} \left| \frac{\hat{q}(\theta) - \alpha_i}{\sigma_{\hat{q}(\theta)}} \right|^{-1}$$

- Here,  $\hat{q}(\theta)$  is GP mean,  $\sigma_{\hat{q}(\theta)}$  is a GP uncertainty,  $\alpha_i$  designates confidence levels, e.g. 68% or 90%
- $a(\theta)$  balances between exploration, i.e. reducing approximation uncertainty, and exploitation, i.e. reaching the extremum.



# Results

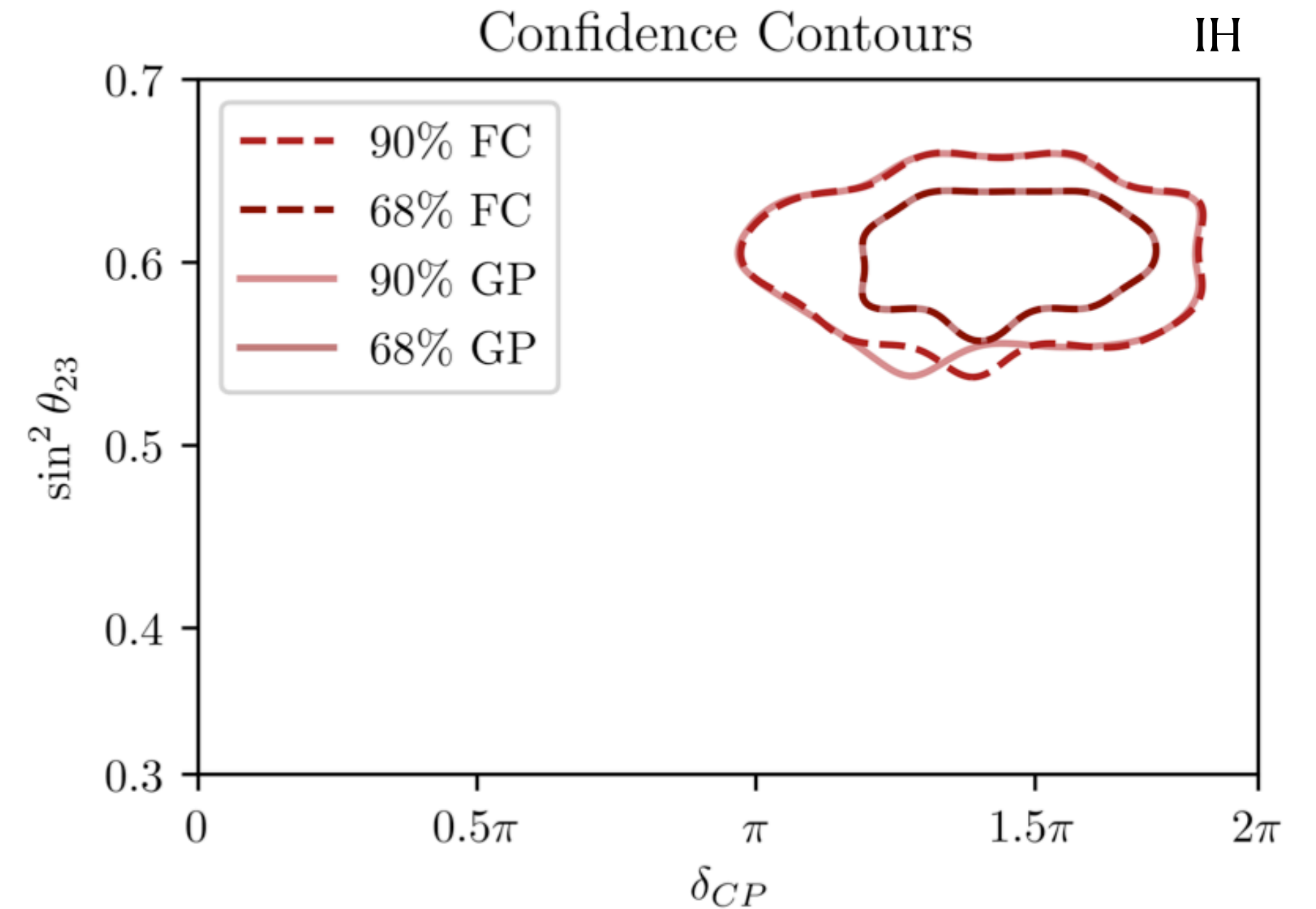
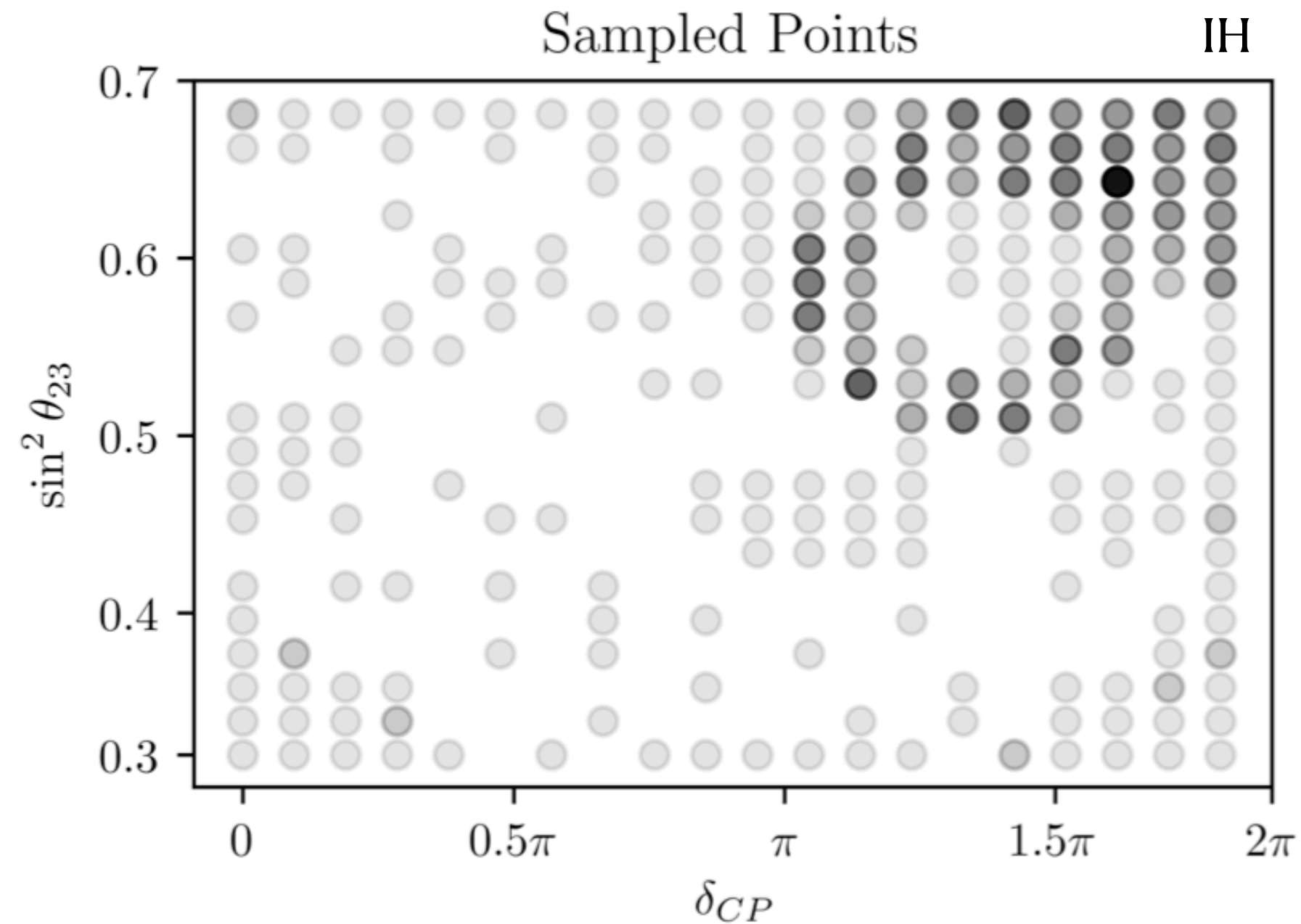
- Oscillation parameters similar to 2018 best estimate from NOvA ( $\theta_{23} = 0.56$ ,  $\Delta m_{32}^2 = 2.44 * 10^{-3} eV^2$ ,  $\delta_{CP} = 1.5\pi$ )
- $\sin^2 \theta_{23} - \delta_{CP}$  68% and 90% CI for NH after 5 iterations.





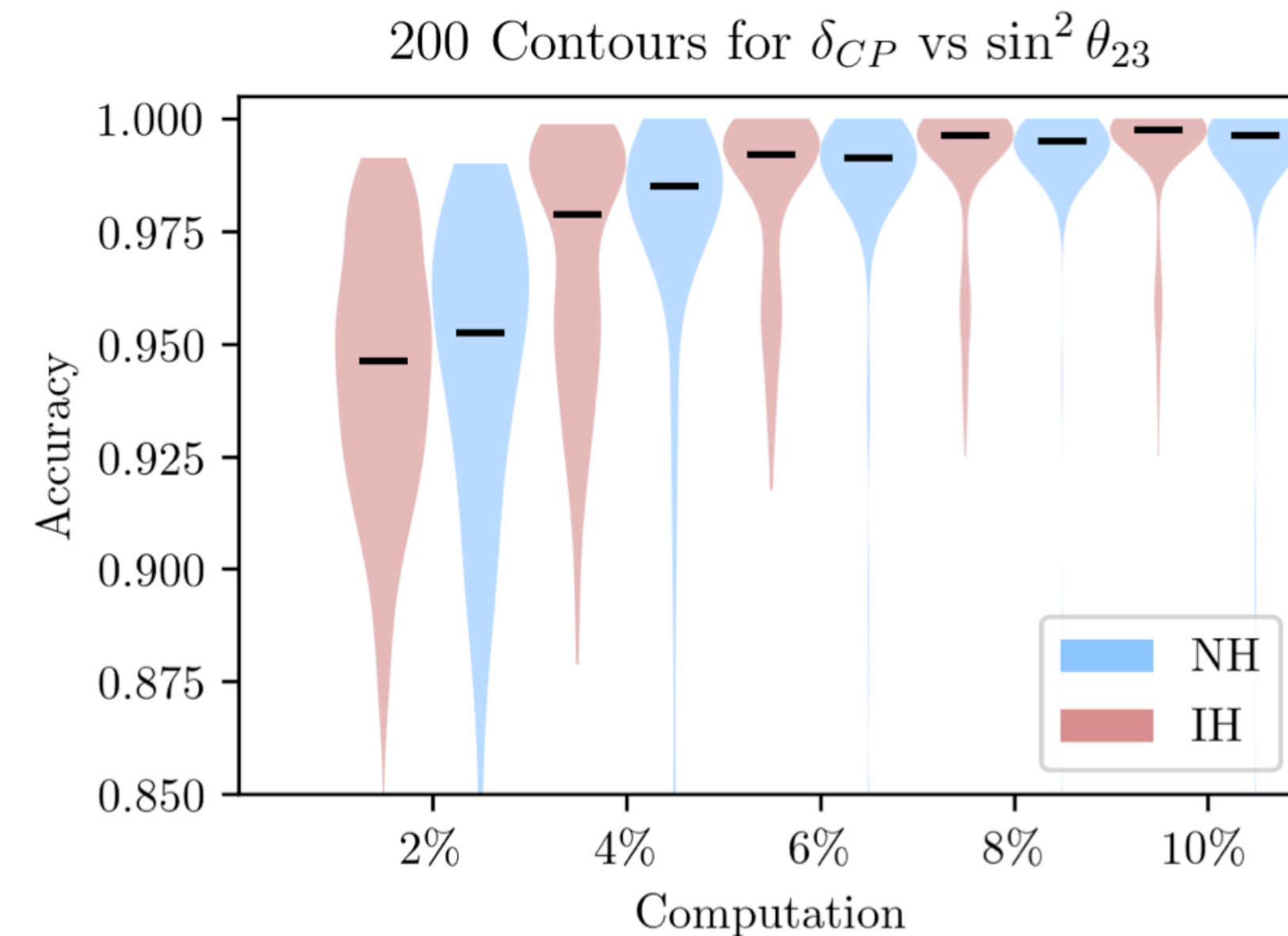
# Results

- Oscillation parameters similar to 2018 best estimate from NOvA ( $\theta_{23} = 0.56$  ,  $\Delta m_{32}^2 = 2.44 * 10^{-3} eV^2$  ,  $\delta_{CP} = 1.5\pi$ )
- $\sin^2 \theta_{23} - \delta_{CP}$  68% and 90% CI for IH after 5 iterations.



# Results

- Use classification accuracy of all grid points, taking FC result as truth, to evaluate performance
- Progress shows the search algorithm converges to the FC value  $\sim 10 \times$  faster for 2D case
- Median Accuracy for for 2D is  $> 99.5\%$  for both NH and IH



# Summary

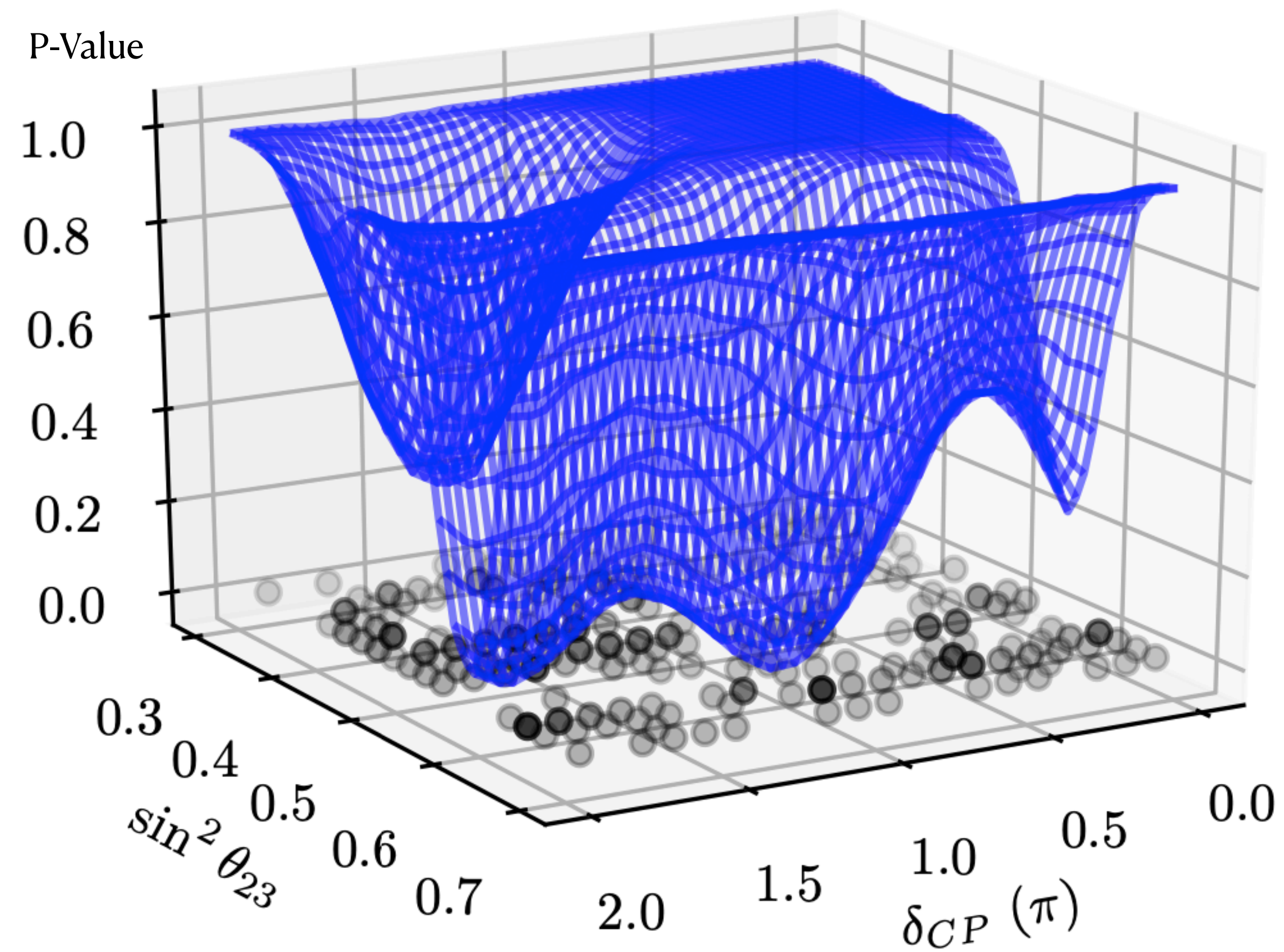
- Neutrino Oscillation experiments provide interesting test case for estimating frequentist confidence intervals.
- LBL experiments typically proceed via Feldman-Cousins
- However, simulating  $\Delta\chi^2$  distribution across multi-dimensional parameter space requires huge computational source
- We've studied Gaussian Process on a toy LBL set-up
- Helps us estimate frequentist contour edges to quite a high accuracy without having to sample the entire parameter space.
- See publication for more details : Phys.Rev.D 101 (2020) 1, 012001
- All code with illustrative notebooks here : <https://github.com/nitish-nayak/ToyNuOscCI>, maintained by Lingge ([linggeli7@gmail.com](mailto:linggeli7@gmail.com)), Nitish ([nayakb@uci.edu](mailto:nayakb@uci.edu)), and Yiwen ([yiwenx7@uci.edu](mailto:yiwenx7@uci.edu))

**Thanks!**

**Backup**

# Results

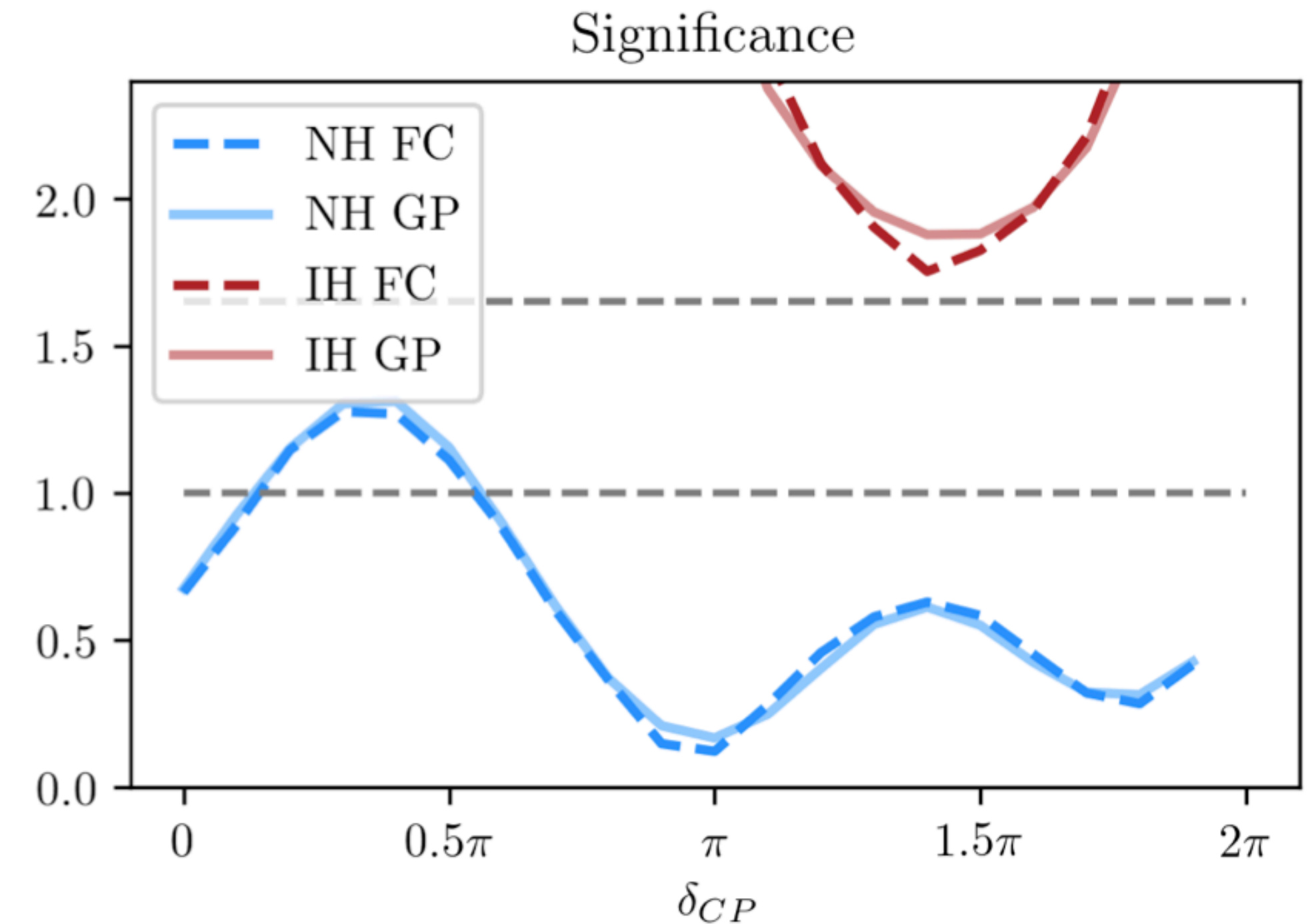
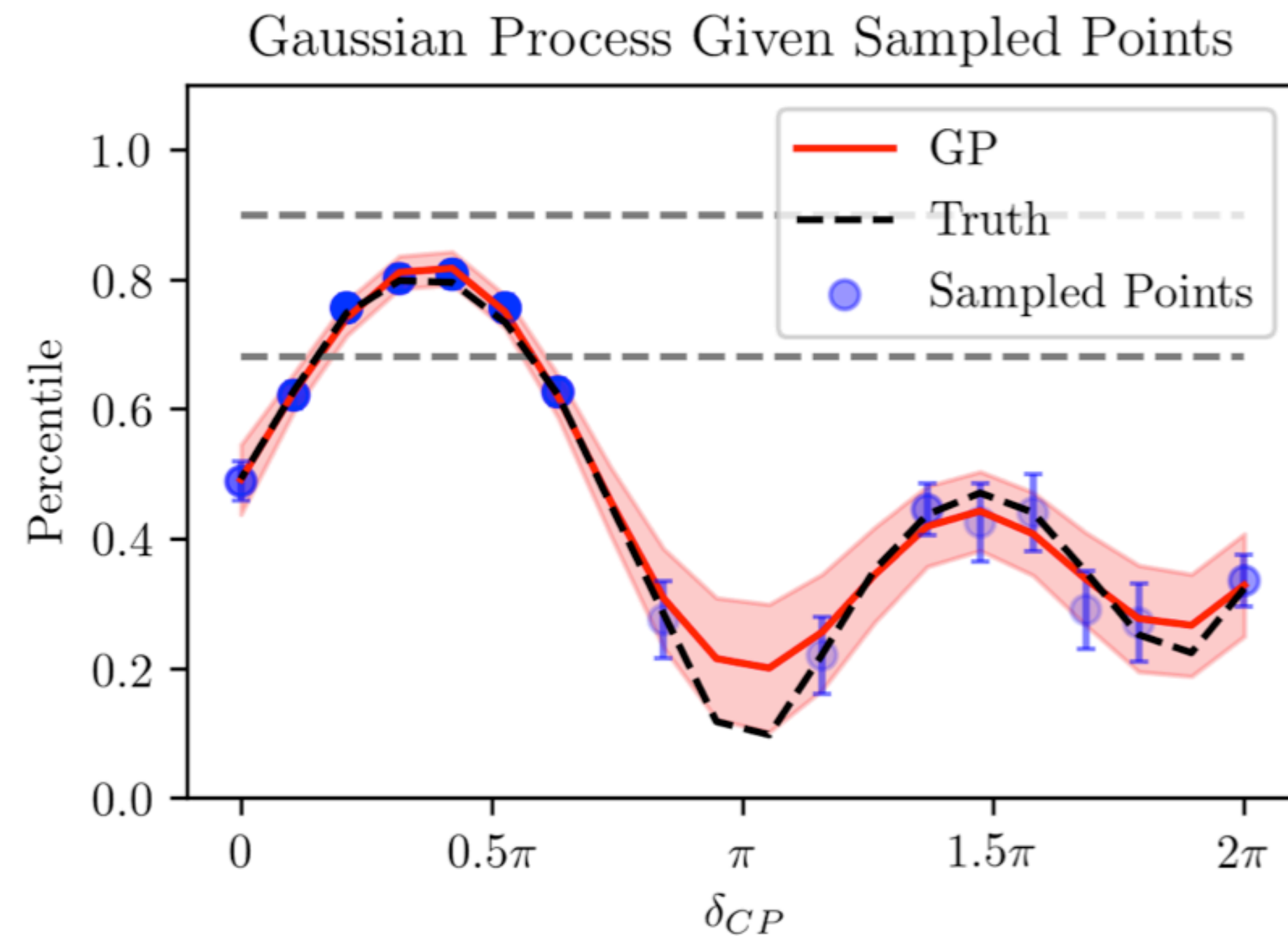
Target Surface





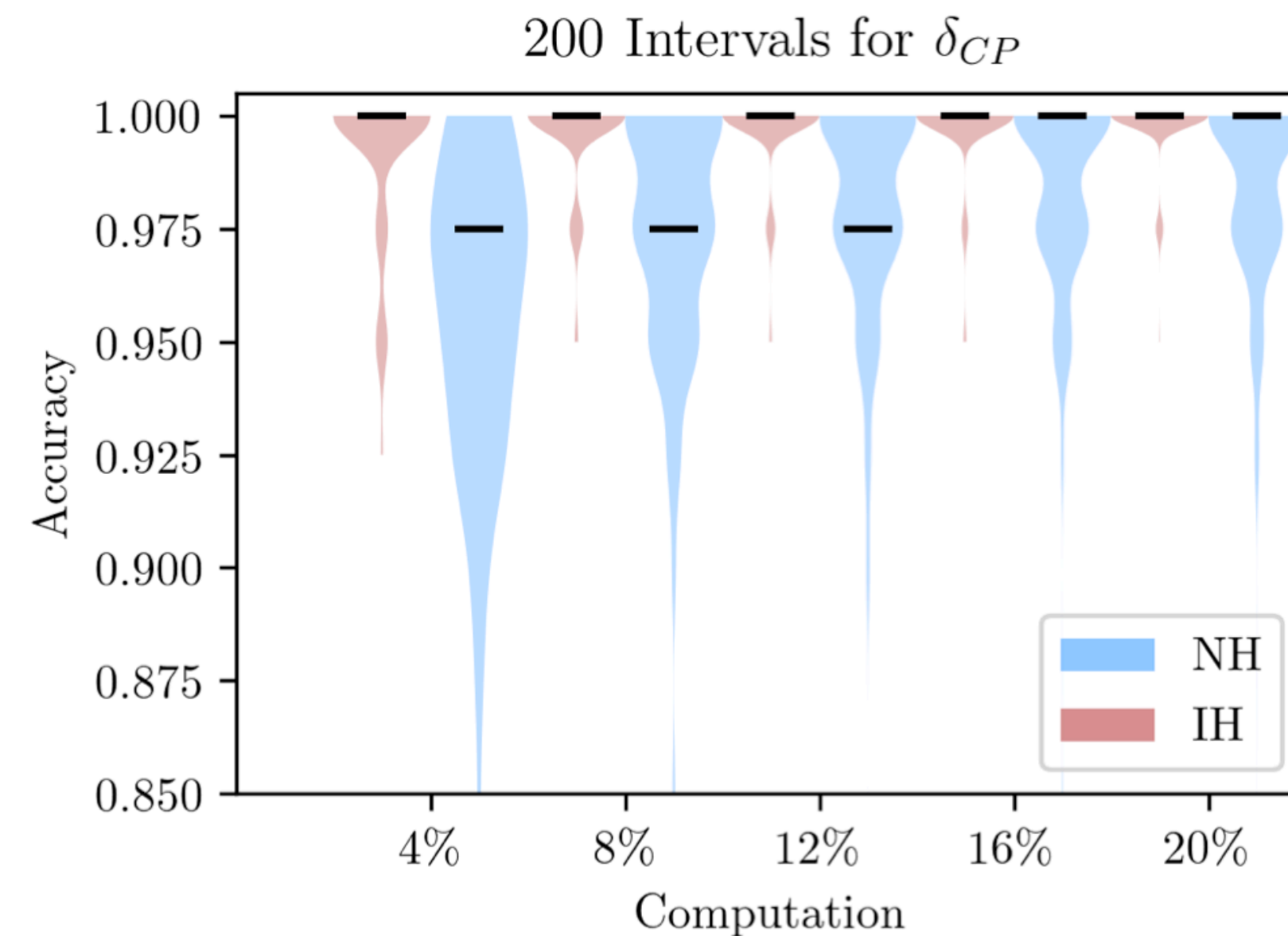
# Results

- Oscillation parameters similar to 2018 best estimate from NOvA ( $\theta_{23} = 0.56$  ,  $\Delta m_{32}^2 = 2.44 * 10^{-3} eV^2$  ,  $\delta_{CP} = 1.5\pi$ )
- Significance of rejecting  $\delta_{CP}$  only after 5 iterations.



# Results

- Use classification accuracy of all grid points, taking FC result as truth, to evaluate performance
- Progress shows the search algorithm converges to the FC value  $\sim 5 \times$  faster for 1D case.
- Median Accuracy for 1D is 100% for both NH and IH



# Pseudo-Code

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**Algorithm 1**  $\mathcal{GP}$  iterative confidence contour finding

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```
for each iteration  $t = 1, 2, \dots$  do
  Propose new points in parameter space  $\arg \max_{\theta} a(\theta)$ 
  for each point  $\theta'$  do
    Simulate likelihood ratio distribution
    for  $k = 1, 2, \dots$  do
      Perform a pseudo experiment
      Maximize the likelihood with respect to  $(\theta, \delta)$ 
      Maximize the likelihood with constraint  $\theta = \theta'$ 
    end for
    Obtain critical value  $c(\theta')$ 
  end for
  Update  $\mathcal{GP}$  approximation  $\hat{c}(\theta)$ 
  Update confidence contours
end for
```

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