# methods for lattice QCD calculations of hadronic observables using stochastic locality

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in collaboration with Mattia Bruno, Anthony Francis, Patrick Fritzsch, Jeremy Green, Maxwell T. Hansen, Antonio Rago

based on

Exploiting stochastic locality in lattice QCD: hadronic observables and their uncertainties arXiv:2307.15674





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#### motivations

#### stochastic locality in QCD:

[Lüscher Lattice 2017]

fields in space-time regions that are far apart fluctuate largely independently

interest in generating large volume lattices

more computational resources available
 advances in communication-avoiding algorithms

[plenary by Boyle]

- strong physics motivations to simulate larger volumes, e.g. reconstructing spectral functions from  $\lim_{\sigma \to 0} \lim_{V \to \infty} \rho_{\sigma,V}$
- the master-field approach is also solution of the topological charge freezing problem, frozen topology bias: O(1/V)  $\ll$  statistical uncertainty:  $O(1/\sqrt{V})$
- $\Rightarrow$  e.g. master-field with  $^{ullet}$  192 $^4$  lattice points  $^{ullet}$  up to pprox 18 fm length  $^{ullet}$   $m_\pi L=25$

[Fritzsch Lattice 2022]

in Bruno, MC, *et al.* arXiv:2307.15674 we investigate methods to exploit stochastic locality in lattice QCD calculations of (hadronic) observables:

in this talk, independent statistics from space-time decorrelation can be used to improve error estimates

#### notation

with one single gauge-field configuration with T and  $L\gg 1/m_\pi$  [Lüscher Lattice 2017, Francis, Fritzsch, Lüscher, Rago 2020] given a local observable  $\mathscr{O}_\alpha(x)$ , the best estimator for the true expectaion value  $\langle \mathscr{O}_\alpha \rangle$ 

$$\langle\!\langle \mathcal{O}_{\alpha} \rangle\!\rangle = \frac{1}{N} \sum_{x} \mathcal{O}_{\alpha}(x)$$

is given by the translation average over a set of points  $x \in \Lambda$ , where  $N = |\Lambda|$  is the number of sample points

the (co)variance of the estimator  $\langle\langle \mathcal{O}_{\alpha} \rangle\rangle$  is

[Lüscher Lattice 2017]

$$\left\langle \left[ \left\langle \left( \mathcal{O}_{\alpha} \right\rangle \right\rangle - \left\langle \mathcal{O}_{\alpha} \right\rangle \right] \left[ \left\langle \left( \mathcal{O}_{\beta} \right\rangle \right\rangle - \left\langle \mathcal{O}_{\beta} \right\rangle \right] \right\rangle = \frac{1}{N^2} \sum_{x,y} \Gamma_{\alpha\beta}(x-y) = \frac{1}{N} \sum_{y} \Gamma_{\alpha\beta}(y)$$

where the correlation function  $\Gamma$  and its sum over  $\gamma$  are

$$\Gamma_{\alpha\beta}(y) = \left\langle \left[ \mathcal{O}_{\alpha}(y) - \left\langle \mathcal{O}_{\alpha} \right\rangle \right] \left[ \mathcal{O}_{\beta}(0) - \left\langle \mathcal{O}_{\beta} \right\rangle \right] \right\rangle, \qquad C_{\alpha\beta} = \sum_{\nu} \Gamma_{\alpha\beta}(y)$$

how to estimate  $\Gamma_{\alpha\beta}$  from one single gauge-field configuration?

# autocorrelation — Wolff's $\Gamma$ method

in a traditional computation with (replicas of) a Monte Carlo chain, for a primary observable  $\mathcal{O}_{\alpha}$ 

[Wolff 2003]

$$\varGamma_{\alpha\beta}(i-j) = \left\langle [\mathcal{O}_{\alpha}^{i} - \left\langle \mathcal{O}_{\alpha} \right\rangle][\mathcal{O}_{\beta}^{j} - \left\langle \mathcal{O}_{\beta} \right\rangle] \right\rangle$$

define an integrated autocorrelation time  $\tau_{\alpha}$ 

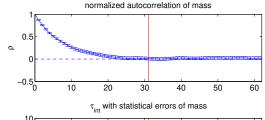
$$\tau_{\alpha} = \frac{1}{2\Gamma_{\alpha\alpha}(0)} \sum_{i=-\infty}^{+\infty} \Gamma_{\alpha\alpha}(i), \quad \text{var}(\mathcal{O}_{\alpha}) = \frac{2\tau_{\alpha}}{N} \Gamma_{\alpha\alpha}(0)$$

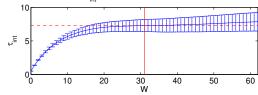
- $\Gamma_{\alpha\alpha}(0)$  is the "naive" variance of  $\mathcal{O}_{\alpha}$
- in practice, extract from i = 1, ..., N configurations

$$\bar{\Gamma}_{\alpha\beta}(t) = \frac{1}{N-t} \sum_{i=1}^{N-t} [\hat{\mathcal{O}}_{\alpha}^{i} - \bar{\mathcal{O}}_{\alpha}] [\hat{\mathcal{O}}_{\beta}^{i+t} - \bar{\mathcal{O}}_{\beta}]$$

- $\sum \bar{\Gamma}(t)$  truncation  $\Rightarrow$  systematic error
- the statistical error of the error is given by Madras-Sokal formula

[Madras, Sokal 1988]





⇒ automatic windowing procedure to balance statistical and systematic uncertainties

# $\Gamma$ method for space-time correlations

straightforward generalization

$$\Gamma_{\alpha\beta}(y) = \left\langle \left[ \mathcal{O}_{\alpha}(y) - \left\langle \mathcal{O}_{\alpha} \right\rangle \right] \left[ \mathcal{O}_{\beta}(0) - \left\langle \mathcal{O}_{\beta} \right\rangle \right] \right\rangle \sim \exp\{-m|y|\}$$

is expected to fall off exponentially with the distance |y| and a mass m

•  $\langle \mathcal{O}_{\alpha} \rangle \neq 0 \Rightarrow$  typically  $m = 2m_{\pi}$  (the energy of the  $0^{++}$  state)

in practice, we replace  $\langle \mathcal{O}_{\alpha} \rangle \rightarrow \langle \langle \mathcal{O}_{\alpha} \rangle \rangle$ 

$$\left\langle\!\left\langle \, \varGamma_{\alpha\beta}(y) \,\right\rangle\!\right\rangle = \frac{1}{N} \, \sum_{x} \, \delta \mathcal{O}_{\alpha}(x+y) \delta \mathcal{O}_{\beta}(x), \qquad \delta \mathcal{O}_{\alpha}(x) = \mathcal{O}_{\alpha}(x) - \left\langle\!\left\langle \, \mathcal{O}_{\alpha} \,\right\rangle\!\right\rangle$$

 $\bullet \ \ \text{a biased estimator: } \left\langle \left\langle \left\langle \left( \varGamma_{\alpha\beta}(y) \right) \right\rangle \right\rangle - \varGamma_{\alpha\beta}(y) = -C_{\alpha\beta}/N$ 

and truncate the sum at a finite summation radius R

$$\langle \langle C_{\alpha\beta}(R) \rangle \rangle = \sum_{|y| < R} \langle \langle \Gamma_{\alpha\beta}(y) \rangle \rangle$$

such that

$$\left\langle \left\langle \left\langle C_{\alpha\beta}(R) \right\rangle \right\rangle \right\rangle = C_{\alpha\beta} \left[ 1 + O(e^{-mR}) + O(1/N) \right]$$

# a higher-dimensional generalization

 $\Lambda$  can be any D-dimensional subspace of space-time, e.g.

$$\begin{split} \Lambda_T &= \{x_0 \mid x_0 \in [0, T-a]\} \\ \Lambda_{TL} &= \{(x_0, x_1) \mid x_0 \in [0, T-a], x_1 \in [0, L-a]\} \\ \Lambda_{L^3} &= \{\vec{x} \mid x_1, x_2, x_3 \in [0, L-a]\} \end{split}$$

or even an irregular subset of randomly sampled points

#### with more than one configuration

e.g. in the case of traditional ensembles of gauge field configurations  $U_i, i = \{1, \dots, N_{MC}\}$ 

$$\left\langle\!\left\langle \Gamma_{\alpha\beta}^{i}(y)\right\rangle\!\right\rangle = \frac{1}{N}\sum_{x}\delta\mathcal{O}_{\alpha}(x+y)\delta\mathcal{O}_{\beta}(x), \qquad \delta\mathcal{O}_{\alpha}^{i}(x) = \mathcal{O}_{\alpha}(x) - \left\langle\!\left\langle \bar{O}_{\alpha}\right\rangle\!\right\rangle$$

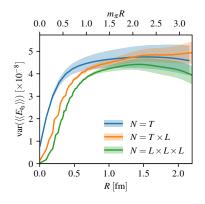
correlations also in MC time ⇒ "five-dimensional" gamma method

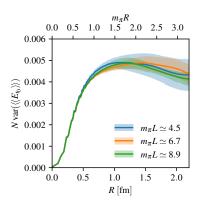
- improve the error estimate if not enough configurations are available
- first explorations along these lines Blum et al. (RBC/UKQCD) 2023

# numerical tests — energy density

of the gauge action with gradient flow at flow-time  $t \approx t_0 = (t \mid t^2 E_t = 0.3)$ 

[Lüscher 2010]





the variance of  $\langle E_{t_0} \rangle$  plateaus for  $R \gtrsim 1.0$  fm

left with different sets of  $\Lambda \Rightarrow$  compatible plateau values, different approach

right and scaling the volume,  $L/a \in \{32, 48, 64\} \Rightarrow m_{\pi}L \in \{4.5, 6.7, 8.9\}$ 

 $\Rightarrow$  perfect scaling of the variance with  $N \propto L^3$ 

### the error of the error

extending Madras-Sokal formula derived in Wolff 2003, see also Madras, Sokal 1988

$$\operatorname{var}\left(\left\langle\!\left\langle C_{\alpha\beta}(R)\right\rangle\!\right\rangle\right) \approx \frac{N(R)}{N} \left[C_{\alpha\alpha}C_{\beta\beta} + C_{\alpha\beta}^{2}\right]$$

where N(R) is the number of points |y| < R,  $\approx \pi^{D/2} (R/b)^D / \Gamma(D/2 + 1)$ 

grows with R
 has to be balanced with the systematic bias

$$\left\langle \left\langle \left\langle C_{\alpha\beta}(R) \right\rangle \right\rangle \right\rangle = C_{\alpha\beta} \left[ 1 + O(e^{-mR}) + O(1/N) \right]$$

as in the  $\Gamma$  method case, we can define an integrated autocorrelation volume

$$\tau_{\alpha} = \frac{C_{\alpha\alpha}}{\Gamma_{\alpha\alpha}(0)} \qquad \Rightarrow \qquad \operatorname{var}(\langle\langle O_{\alpha} \rangle\rangle) = \tau_{\alpha} \Gamma_{\alpha\alpha}(0)/N$$

- dimension D dependent definition!
  - estimator as a function of R

$$\tau_{\alpha}(R) = \frac{1}{\Gamma_{\alpha\alpha}(0)} \left\langle \left\langle C_{\alpha\beta}(R) \right\rangle \right\rangle = \frac{1}{\Gamma_{\alpha\alpha}(0)} \sum_{|y| < R} \left\langle \left\langle \Gamma_{\alpha\alpha}(y) \right\rangle \right\rangle$$

# hadronic observables

the mesonic two-point function projected to zero momentum,  $\Gamma,\Gamma'\in\{\gamma_5,\gamma_\mu,\gamma_5\gamma_\mu,\dots\}$ 

$$\tilde{C}(x_0 - y_0, \vec{x}) = -a^3 \sum_{\vec{y}} \text{Re Tr} [\Gamma D^{-1}(y, x) \Gamma' D^{-1}(x, y)]$$

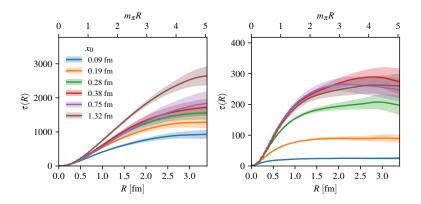
estimated using random 3d-volume source  $\Rightarrow \vec{x} \in \varLambda_{L^3}$  has an error given by

$$\left\langle \left[ \left\langle \left( \tilde{C}(t) \right) \right\rangle - \left\langle \tilde{C}(t) \right\rangle \right]^{2}(R) \right\rangle = \frac{1}{L^{3}} \left[ \sum_{|y| \leq R} \left\langle \left( \tilde{C}(t; \vec{x}) \tilde{C}(t; 0) \right) \right\rangle_{c} + O(e^{-mR}) + O(L^{-3}) \right]$$

where each source-sink separation t defines a different observable

# scaling with source-sink separation

left: pseudoscalar  $\tilde{C}_{PP}$ , right: vector  $\tilde{C}_{VV}$ , each source-sink separation defines a different observable

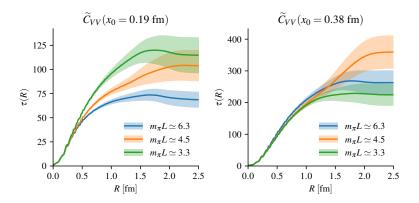


#### important: saturation is a function of the statistical precision!

- at large source-sink separations, the noise hides space-time correlations
- just as in the case of autocorrelations in Monte Carlo time

# scaling with pion mass

as a function of  $m_\pi \in \{215, 293, 410\}$  MeV with L=32a, for  $\tau(R)$  of the vector correlator  $\tilde{\mathcal{C}}_{VV}$ 



left: at very short distance  $x_0 \approx 0.19\,\mathrm{fm}$ : lighter pions  $\Rightarrow$  larger  $\tau(R)$ , plateau at larger R right: at  $x_0 \approx 0.38\,\mathrm{fm}$  or larger: noisier data  $\Rightarrow \tau(R)$  approchaes are comparable

# extrapolation to large volumes

our investigation on moderately-large volumes up to  $96 \times 64^3$  /  $m_\pi L \approx 8.9$ 

 $\Rightarrow$  fit of the variance of the pion correlator at source-sink separation  $x_0=4a$ 

⇒ estimate the error of the same observable on a different volume

e.g. we project a 0.5% error on the pseudoscalar correlator from a single configuration with T=L=192a

preliminary exploration: numerical tests a single  $192^4$  /  $m_\pi L \approx 25$  master-field configuration (same lattice spacing  $a\approx 0.095$  fm but slightly different  $m_\pi\approx 270\,\text{MeV!}$ ) stochastic 3d-volume sources at two source times:  $x_0=0$  and 96a

[Fritzsch Lattice 2022]

$$\tilde{C}_{PP}(x_0=4a)=0.061\,01(23) \quad \text{and} \quad am_\pi=0.126\,32(27)$$

a 0.38% and 0.22% error respectively, obtained using pyobs

#### conclusions

- ullet very general method, Wollf's  $\Gamma$  method is the 1d version of this
- already used in recent studies to improve error estimates with few configurations

[Blum et al. (RBC/UKQCD) 2023]

- large-separation correlators have a larger footprint in space
   ⇒ estimation of the integrated correlation volume depends on the statistics
- the whole procedure can be automated, see e.g. M. Bruno's pyobs Python package

https://mbruno46.github.io/pyobs/

- results for  $\Gamma_{\alpha\beta}$  on a intermediate-volume lattices
  - ⇒ can inform (at fixed lattice spacing) and help plan large-volume / master-field simulation

details in the recent paper: Bruno, MC, et al. arXiv:2307.15674

## position-space correlators

studied on the same intermediate volumes

[MC et al. Lattice 2021; Bruno, MC, et al. arXiv:2307.15674]

• and preliminary results on 96<sup>4</sup>, 192<sup>4</sup> master fields

[MC et al. Lattice 2022]

ullet using blocking; or  $\Gamma$  method with a sufficiently high density of points

long-T approach as a solution of topological charge freezing

[Francis et al. Lattice 2022; Bruno, MC, et al. arXiv:2307.15674]

# for your attention!

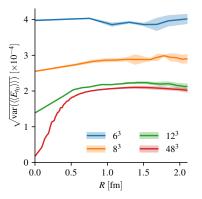
thanks

questions?

# backup slides

## saturation of the error

scaling the density of the points on which the observable  $E_{t_0}$  is computed



- red line at  $48^3$  is each point on the time slice, i.e.  $\Lambda_{L^3}$
- green, orange, blue: artificial subsampling
- the error increases for a very sparse grid
- the error saturates as a the grid get denser:

  orange:  $N=12^3$  point samples on a grid with spacing 4a  $\Rightarrow$  same error as the whole time slice (red)

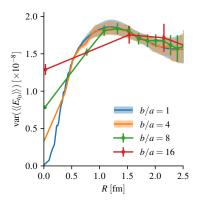
⇒ possible strategy for hadronic observables: evaluate (a) stochastic volume sources vs. (b) a (more or less dense) grid of source points

source points can also be randomly distributed

# blocking

of the observable  $\mathcal{O}_{\alpha}$  over small blocks of size  $b^4$ 

$$\mathcal{O}_{B\alpha}(u) = \frac{N_B}{N} \sum_{x \in \mathsf{block}} \mathcal{O}_{\alpha}(x)$$



- this can significantly reduce the cpu and memory footprint of the  $\Gamma$  method, in case of very large volumens
- moderate blocking does not compromise the error of the error
- the scaling of  $\Gamma_{\alpha\beta}$  is non-trivial,  $\mathrm{var}(\mathcal{O}_{\alpha}) = \tau_{\alpha}\Gamma_{\alpha\alpha}/N$   $\Rightarrow \tau_{\alpha}(R)$  depends on the blocking possible solution: define  $\tau_{\alpha}$  with the non-blocked  $\Gamma_{\alpha\alpha}(0)$

e.g.  $192^4$  lattice points  $\Rightarrow 48^4$  blocks of length 4a

# truncated 3d sums

$$\tilde{C}^{\text{cut}}(t, r_{\text{max}}) = \int d^3 \vec{x} \, \theta(r_{\text{max}} - \left| \vec{x} \right|) C(\vec{x}, t)$$

- explored in e.g. Liu, Liang, Yang 2018
- $\bullet$  larger t require larger  $r_{\rm max}$  to reach the same  $\tilde{C}^{\rm cut}/\tilde{C}$
- the pion correlator (top) is especially slow, does not saturate!
- the nucleon one (bottom) saturates at  $r_{\rm max} \approx 20a$ , but also the error saturates

⇒ no statistical benefit

