

Faster Monte Carlo via Low Discrepancy Sampling

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Thank you for your invitation and hospitality,

Slides at speakerdeck.com/fjhickernell/argonne2023maytalk

Jupyter notebook with computations and figures [here](#) Visit us at qmcpy.org

Argonne Seminar, revised Wednesday 24th May, 2023





Family of Physicists

- Fred S. Hickernell (father), PhD **Physics**, IEEE Fellow, Motorola researcher
- Robert K. Hickernell (brother), PhD **Physics**, NIST Division Director
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- Myself, BA mathematics and physics who became an **applied mathematician**



Message

Problems in Bayesian inference, uncertainty quantification, quantitative finance, (high energy) physics¹, etc. require **numerically** evaluating

$$\underbrace{\mathbb{E}[f(\mathbf{X})]}_{\text{expectation}} = \underbrace{\int_{\Omega} f(\mathbf{x}) \varrho(\mathbf{x}) \, d\mathbf{x}}_{\text{integral}}, \quad \mathbf{X} \sim \varrho$$

Other more complicated problems include computing quantiles or marginal distributions.

¹M. R. Blaszkiewicz (Feb. 2022). “Methods to optimize rare-event Monte Carlo reliability simulations for Large Hadron Collider Protection Systems”. MA thesis. University of Amsterdam. URL: <https://cds.cern.ch/record/2808520>; A. Courtoy, J. Huston, et al. (2023). “Parton distributions need representative sampling”. In: *Phys. Rev. D* 107.034008. URL: <https://journals.aps.org/prd/pdf/10.1103/PhysRevD.107.034008>; D. Everett, W. Ke, et al. (May 2021). “Multisystem Bayesian constraints on the transport coefficients of QCD matter”. In: *Phys. Rev. C* 103.5. DOI: 10.1103/physrevc.103.054904; D. Liyanage, Y. Ji, et al. (Mar. 2022). “Efficient emulation of relativistic heavy ion collisions with transfer learning”. In: *Phys. Rev. C* 105.3. DOI: 10.1103/physrevc.105.034910.



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There is value in

- **low discrepancy** sampling, aka **quasi-Monte Carlo** methods
- data-driven **error bounds**
- **flattening** and **reducing the effective dimension** of the integrand
- quality quasi-Monte Carlo **software** like **qmcpy**¹
- physicists and quasi-Monte Carlo theorists **collaborating** more

¹S.-C. T. Choi, F. J. H., et al. (2023). *QMCPy: A quasi-Monte Carlo Python Library (versions 1–1.4)*. DOI: 10.5281/zenodo.3964489. URL: <https://qmcsoftware.github.io/QMCSoftware/>.



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Caveat: I know little about Markov Chain Monte Carlo (MCMC); limited success combining with quasi-Monte Carlo².

²S. Chen, J. Dick, and A. B. Owen (2011). "Consistency of Markov Chain Quasi-Monte Carlo on Continuous State Spaces". In: *Ann. Stat.* 9, pp. 673–701; Art B. Owen and Seth D. Tribble (2005). "A quasi-Monte Carlo Metropolis algorithm". In: *Proc. Natl. Acad. Sci.* 102, pp. 8844–8849. 3/25

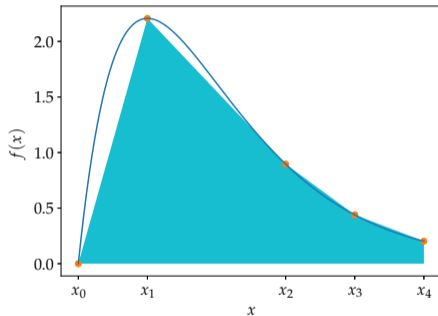


Trapezoidal rule

Trio identity³ — the error of approximating an expectation or integral expressed as the product of **three** quantities

$$\int_0^1 f(x) dx - \sum_{i=0}^n w_i f(x_i) = 0.112 \quad w_i = \frac{x_{i+1} - x_{i-1}}{2}$$

$$= \underbrace{\frac{\int_0^1 f(x) dx - \sum_{i=0}^n w_i f(x_i)}{\frac{\max_i (x_{i+1} - x_i)^2}{8}}}_{\text{misfortune } 0.151} \underbrace{\frac{\max_i (x_{i+1} - x_i)^2}{8}}_{\text{sampling deficit } 0.020} \underbrace{\int_0^1 |f''(x)| dx}_{\text{roughness } 37.3}$$



³F. J. H. (2018). "The Trio Identity for Quasi-Monte Carlo Error Analysis". In: *Monte Carlo and Quasi-Monte Carlo Methods: MCQMC, Stanford, USA, August 2016*. Ed. by P. Glynn and A. Owen. Springer Proceedings in Mathematics and Statistics. Springer-Verlag, Berlin, pp. 3–27. doi: 10.1007/978-3-319-91436-7; A. Courtoy, J. Huston, et al. (2023). "Parton distributions need representative sampling". In: *Phys. Rev. D* 107.034008. URL: <https://journals.aps.org/prd/pdf/10.1103/PhysRevD.107.034008>; X. Meng (2018). "Statistical Paradises and Paradoxes in Big Data (I): Law of Lagne Populations, Big Data Paradox, and 2016 US Presidential Election". In: *Ann. Appl. Stat.* 12, pp. 685–726. 4/25



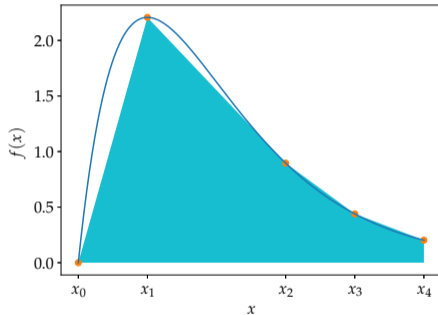
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- Confounding is **between -1 and 1** (Brass and Petras 2011, (7.15))
- Discrepancy depends upon **sample** only
- Variation depends on **integrand** only, semi-norm



³F. J. H. (2018). "The Trio Identity for Quasi-Monte Carlo Error Analysis". In: *Monte Carlo and Quasi-Monte Carlo Methods: MCQMC, Stanford, USA, August 2016*. Ed. by P. Glynn and A. Owen. Springer Proceedings in Mathematics and Statistics. Springer-Verlag, Berlin, pp. 3–27. doi: 10.1007/978-3-319-91436-7; A. Courtoy, J. Huston, et al. (2023). "Parton distributions need representative sampling". In: *Phys. Rev. D* 107.034008. URL: <https://journals.aps.org/prd/pdf/10.1103/PhysRevD.107.034008>; X. Meng (2018). "Statistical Paradises and Paradoxes in Big Data (I): Law of Lagne Populations, Big Data Paradox, and 2016 US Presidential Election". In: *Ann. Appl. Stat.* 12, pp. 685–726.



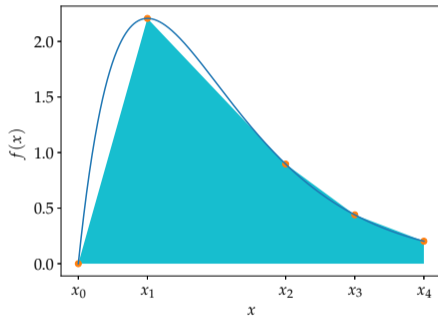
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- Confounding is **between -1 and 1** (Brass and Petras 2011, (7.15))
- Discrepancy reduced via **clever** or more sampling
- Variation value unknown, reduced via **transformations**



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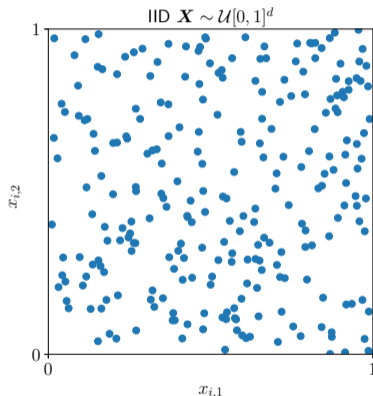
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Independent and Identically Distributed (IID) Monte Carlo

$$\begin{aligned}
 & \underbrace{\mathbb{E}[f(\mathbf{X})], \mathbf{X} \sim \varrho}_{\int_{\mathbb{R}^d} f(\mathbf{x}) \varrho(\mathbf{x}) \, d\mathbf{x}} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i) \quad \mathbf{x}_i \stackrel{\text{iid}}{\sim} \varrho \\
 &= \underbrace{\frac{\int_{\mathbb{R}^d} f(\mathbf{x}) \varrho(\mathbf{x}) \, d\mathbf{x} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i)}{\frac{1}{\sqrt{n}} \text{std}[f(\mathbf{X})]}}_{\text{confounding}} \underbrace{\frac{1}{\sqrt{n}}}_{\text{discrepancy}} \underbrace{\text{std}(f(\mathbf{X}))}_{\text{variation}}
 \end{aligned}$$

- $\text{RMS}[\text{confounding}] = 1$, but confounding could be $\mathcal{O}(\sqrt{n})$
- Discrepancy depends on sample size only
- Variation reduced through transformations

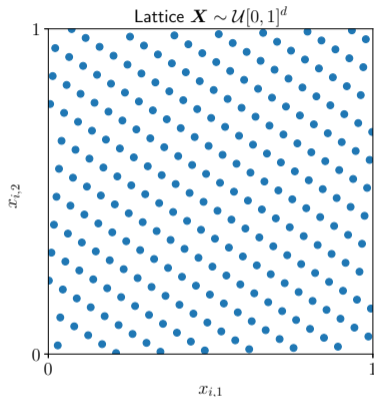




Low Discrepancy Sampling aka Quasi-Monte Carlo⁴

$$\underbrace{\mathbb{E}[f(\mathbf{X})], \mathbf{X} \sim \mathcal{U}[0,1]^d}_{\int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x}} - \frac{1}{n} \sum_{i=1}^n f(\mathbf{x}_i) = \text{CNF}(f, \{\mathbf{x}\}_{i=1}^n) \text{DSC}(\{\mathbf{x}\}_{i=1}^n) \text{VAR}(f)$$

$$\begin{aligned} \text{VAR}^2(f) &= \int_{[0,1]} \left[\frac{\partial f}{\partial x_1}(x_1, 0.5, \dots, 0.5) \right]^2 dx_1 + \dots \\ &+ \int_{[0,1]^2} \left[\frac{\partial^2 f}{\partial x_1 \partial x_2}(x_1, x_2, 0.5, \dots, 0.5) \right]^2 dx_1 dx_2 + \dots \\ &+ \int_{[0,1]^d} \left[\frac{\partial^d f}{\partial x_1 \dots \partial x_d}(\mathbf{x}) \right]^2 d\mathbf{x} \end{aligned}$$



VAR is a semi-norm, more smoothness than std, value generally unknown, reduced through transformations

⁴F. J. H. (1998). "A Generalized Discrepancy and Quadrature Error Bound". In: *Math. Comp.* 67, pp. 299–322. doi: 10.1090/S0025-5718-98-00894-1.



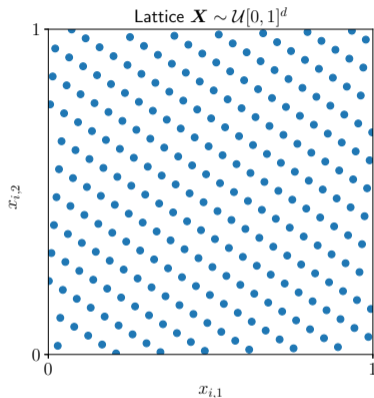
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$$\text{DSC}^2(\{\mathbf{x}_i\}_{i=1}^n) = \left(\frac{13}{12}\right)^2$$

$$- \frac{2}{n} \sum_{i=1}^n \prod_{j=1}^d (1 + 0.5 |x_{ij} - 0.5| - 0.5(x_{ij} - 0.5)^2)$$

$$+ \frac{1}{n^2} \sum_{i,k=1}^n [1 + 0.5 |x_{ij} - 0.5| + 0.5 |x_{kj} - 0.5| - 0.5 |x_{ij} - x_{kj}|]$$



$$\text{DSC}(\{\mathbf{x}_i\}_{i=1}^n) = \mathcal{O}(n^{-1+\delta})$$

DSC is the norm of the error functional, value known with $\mathcal{O}(dn^2)$ operations

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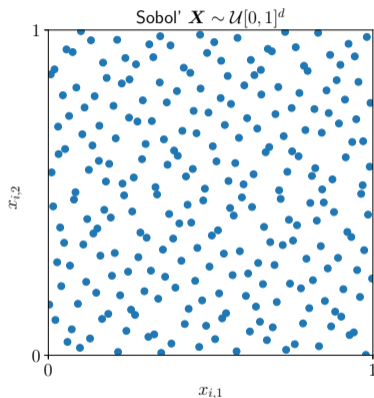
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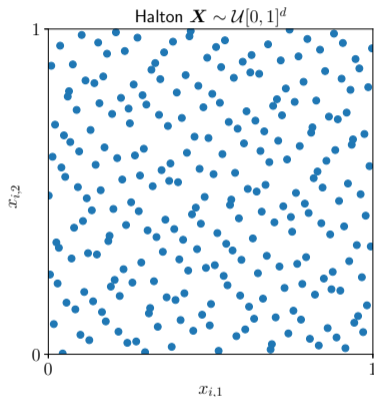
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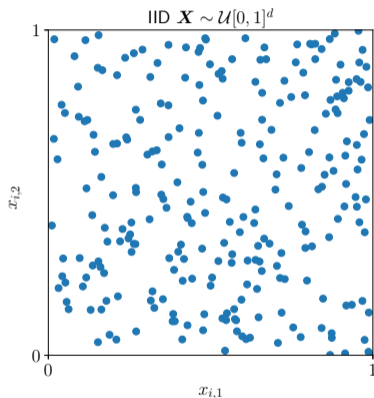
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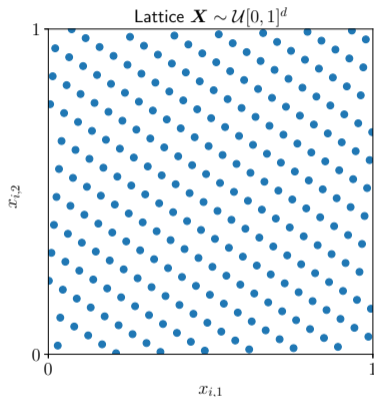
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$$\text{CNF}(f, \{\mathbf{x}_i\}_{i=1}^n) = \frac{\int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x} - \frac{1}{n} \sum_{i=1}^n f(\mathbf{x}_i)}{\text{DSC}(\{\mathbf{x}_i\}_{i=1}^n) \text{VAR}(f)} \quad \text{between } \pm 1$$



$$\text{DSC}(\{\mathbf{x}_i\}_{i=1}^n) = \mathcal{O}(n^{-1+\delta})$$

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Uncertainty in a Cantilevered Beam⁵

$u(x) = g(\mathbf{Z}, x) =$ beam deflection

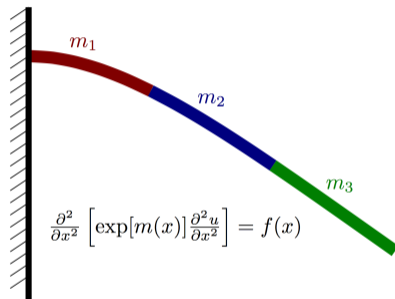
= solution of a differential equation boundary value problem

$\mathbf{Z} \sim \mathcal{U}[1, 1.2]^3$ defines uncertainty in Young's modulus

$x =$ position

$$\mu(x) = \mathbb{E}[g(\mathbf{Z}, x)] = \int_{[0,1]^3} g(\mathbf{z}, x) \, d\mathbf{z} \approx \frac{1}{n} \sum_{i=1}^n g(\mathbf{Z}_i, x)$$

$$\mu(\text{end}) = 1037$$



⁵M. Parno and L. Seelinger (2022). *Uncertainty propagation of material properties of a cantilevered beam*. URL:



IID vs. Low Discrepancy for the Cantilevered Beam via QMCPy⁶

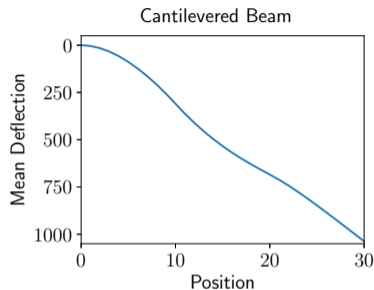
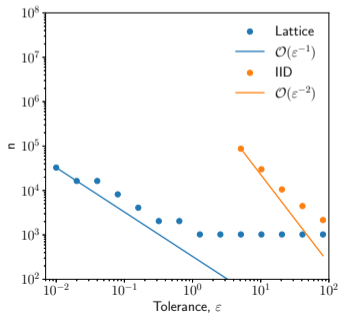
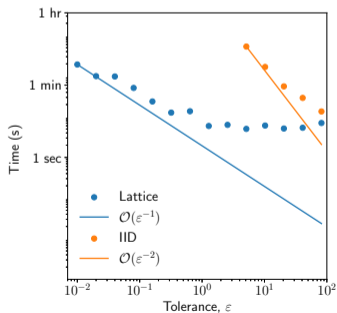
$g(\mathbf{Z}, x) =$ beam deflection

$\mathbf{Z} \sim \mathcal{U}[1, 1.2]^3$ defines uncertainty

$x =$ position

$\mu(x) = \mathbb{E}[g(\mathbf{Z}, x)]$

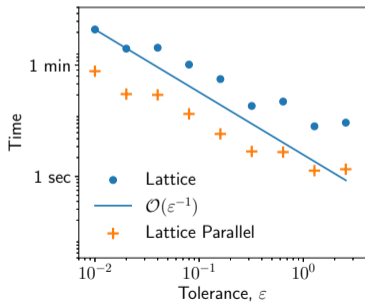
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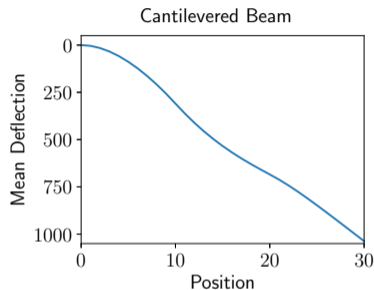
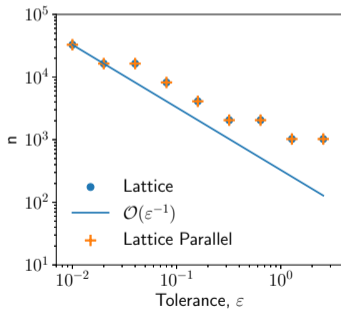
⁶S.-C. T. Choi, F. J. H., et al. (2023). QMCPy: A quasi-Monte Carlo Python Library (versions 1–1.4). DOI: 10.5281/zenodo.3964489. URL: <https://qmcsoftware.github.io/QMCSoftware/>.

IID vs. Low Discrepancy for the Cantilevered Beam via QMCPy⁶ $g(\mathbf{Z}, x) = \text{beam deflection}$

$$\mu(x) = \mathbb{E}[g(\mathbf{Z}, x)]$$



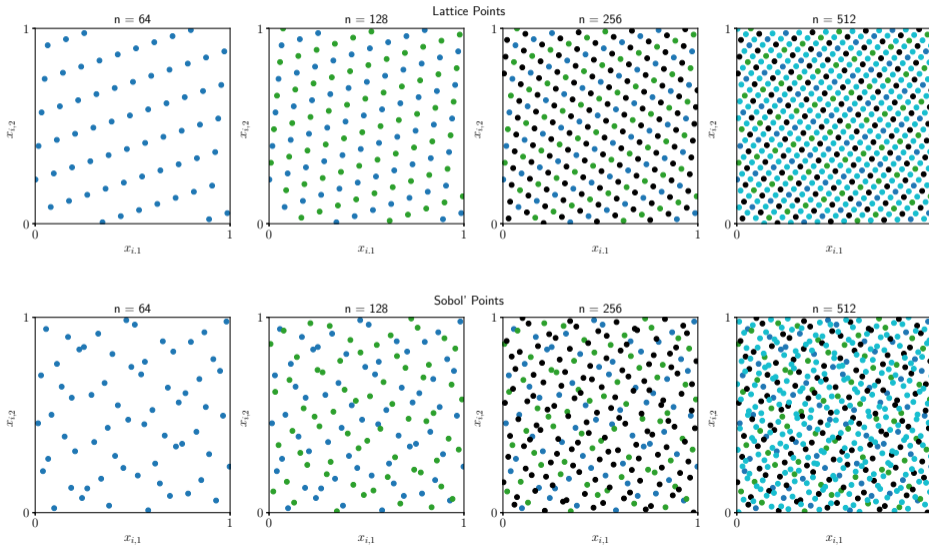
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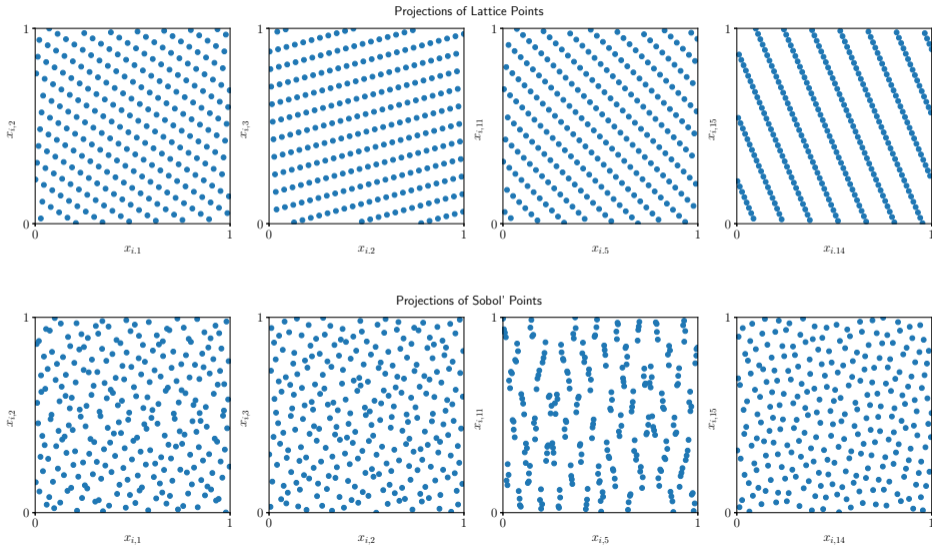


Low Discrepancy Points Fill Space



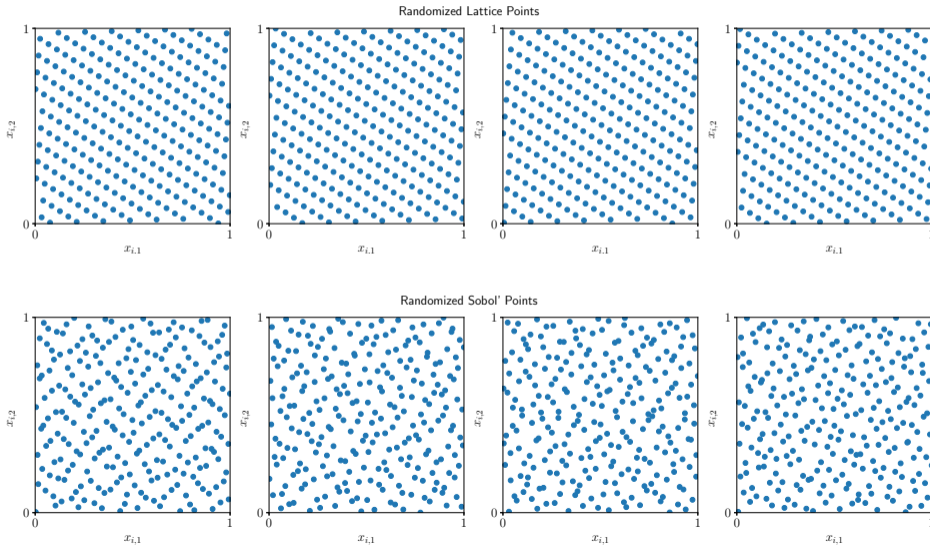


Low Discrepancy Points Look “Good” in All Coordinate Projections





Low Discrepancy Points Can Be Randomized





Lessons from the Trio Identity

$$\mathbb{E}[f(\mathbf{X})], \mathbf{X} \sim \mathcal{U}[0,1]^d$$

$$\int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i) = \text{CNF}(f, \{\mathbf{x}\}_{i=1}^n) \text{ DSC}(\{\mathbf{x}\}_{i=1}^n) \text{ VAR}(f)$$

- Use **low discrepancy**⁷ instead of IID sampling for performance gains

⁷J. Dick, P. Kritzer, and F. Pillichshammer (2022). *Lattice Rules: Numerical Integration, Approximation, and Discrepancy*. Springer Series in Computational Mathematics. Springer Cham. doi: <https://doi.org/10.1007/978-3-031-09951-9>; J. Dick, F. Kuo, and I. H. Sloan (2013). “High dimensional integration — the Quasi-Monte Carlo way”. In: *Acta Numer.* 22, pp. 133–288. doi: 10.1017/S0962492913000044; J. Dick and F. Pillichshammer (2010). *Digital Nets and Sequences: Discrepancy Theory and Quasi-Monte Carlo Integration*. Cambridge: Cambridge University Press.



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- Use **low discrepancy**⁷ instead of IID sampling for performance gains
- **Randomize** when possible so avoid bad $\text{CNF}(f, \{\mathbf{x}\}_{i=1}^n)$.

⁷J. Dick, P. Kritzer, and F. Pillichshammer (2022). *Lattice Rules: Numerical Integration, Approximation, and Discrepancy*. Springer Series in Computational Mathematics. Springer Cham. doi: <https://doi.org/10.1007/978-3-031-09951-9>; J. Dick, F. Kuo, and I. H. Sloan (2013). “High dimensional integration — the Quasi-Monte Carlo way”. In: *Acta Numer.* 22, pp. 133–288. doi: 10.1017/S0962492913000044; J. Dick and F. Pillichshammer (2010). *Digital Nets and Sequences: Discrepancy Theory and Quasi-Monte Carlo Integration*. Cambridge: Cambridge University Press.



Lessons from the Trio Identity

$$\underbrace{\mathbb{E}[f(\mathbf{X})], \mathbf{X} \sim \mathcal{U}[0,1]^d}_{\int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x}} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i) = \text{CNF}(f, \{\mathbf{x}\}_{i=1}^n) \text{DSC}(\{\mathbf{x}\}_{i=1}^n) \text{VAR}(f)$$

- Use **low discrepancy**⁷ instead of IID sampling for performance gains
- **Randomize** when possible so avoid bad $\text{CNF}(f, \{\mathbf{x}\}_{i=1}^n)$. Scrambled Sobol' often beats unscrambled Sobol' in order of convergence. Randomizing moves points off the boundaries.
- Error decay rate, $\text{DSC}(\{\mathbf{x}\}_{i=1}^n)$, **limited** by the assumptions on f implicit in the choice of VAR
- **Deterministic** trio identities constrain $-1 \leq \text{CNF}(f, \{\mathbf{x}\}_{i=1}^n) \leq 1$ but may be **pessimistic**
- If $\text{CNF}(f, \{\mathbf{x}\}_{i=1}^n)$ is consistently small, look for a **better choice** of VAR and DSC

⁷J. Dick, P. Kritzer, and F. Pillichshammer (2022). *Lattice Rules: Numerical Integration, Approximation, and Discrepancy*. Springer Series in Computational Mathematics. Springer Cham. doi: <https://doi.org/10.1007/978-3-031-09951-9>; J. Dick, F. Kuo, and I. H. Sloan (2013). "High dimensional integration — the Quasi-Monte Carlo way". In: *Acta Numer.* 22, pp. 133–288. doi: 10.1017/S0962492913000044; J. Dick and F. Pillichshammer (2010). *Digital Nets and Sequences: Discrepancy Theory and Quasi-Monte Carlo Integration*. Cambridge: Cambridge University Press.



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- If $\text{CNF}(f, \{\mathbf{x}\}_{i=1}^n)$ is consistently small, look for a **better choice** of VAR and DSC
- **Theory** explains how well an algorithm works; practice can help sharpen theory

⁷J. Dick, P. Kritzer, and F. Pillichshammer (2022). *Lattice Rules: Numerical Integration, Approximation, and Discrepancy*. Springer Series in Computational Mathematics. Springer Cham. doi: <https://doi.org/10.1007/978-3-031-09951-9>; J. Dick, F. Kuo, and I. H. Sloan (2013). "High dimensional integration — the Quasi-Monte Carlo way". In: *Acta Numer.* 22, pp. 133–288. doi: 10.1017/S0962492913000044; J. Dick and F. Pillichshammer (2010). *Digital Nets and Sequences: Discrepancy Theory and Quasi-Monte Carlo Integration*. Cambridge: Cambridge University Press.



Estimating or Bounding the Error from Simulation Data

$$\overbrace{\int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x}}^{\mathbb{E}[f(X)], X \sim \mathcal{U}[0,1]^d} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i) = \text{CNF}(f, \{\mathbf{x}\}_{i=1}^n) \text{DSC}(\{\mathbf{x}\}_{i=1}^n) \text{VAR}(f)$$

When to **stop** sampling because the error is small enough?

- $\text{VAR}(f)$ is **impractical** to bound
- $\text{DSC}(\{\mathbf{x}\}_{i=1}^n)$ is **expensive** to calculate



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Many do replications

$$\left| \int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i) \right|^2 \approx \frac{\text{fudge}^2}{R} \sum_{r=1}^R \left(\frac{1}{Rn} \sum_{q,i=1}^{R,n} f(\mathbf{x}_i^{(q)}) - \frac{1}{n} \sum_{i=1}^n f(\mathbf{x}_i^{(r)}) \right)^2$$

where $\{\mathbf{x}_i^{(1)}\}_{i=1}^n, \dots, \{\mathbf{x}_i^{(R)}\}_{i=1}^n$ are randomizations of a low discrepancy sequence.



Estimating or Bounding the Error from Simulation Data

For $\left\{ \begin{array}{l} \text{lattice} \\ \text{Sobol'} \end{array} \right\}$ or sampling

$$\left| \int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i) \right| \leq \sum_{\mathbf{k} \in \text{dual lattice}} |\widehat{f}(\mathbf{k})|$$

where $\widehat{f}(\mathbf{k})$ are the Fourier $\left\{ \begin{array}{l} \text{complex exponential} \\ \text{Walsh} \end{array} \right\}$ coefficients of f and the **dual lattice** consists of wave numbers of $\left\{ \begin{array}{l} \text{complex exponential} \\ \text{Walsh} \end{array} \right\}$ bases that are perfectly aliased with 1.

Rigorous data-driven error bounds based on sums of discrete Fourier $\left\{ \begin{array}{l} \text{complex exponential} \\ \text{Walsh} \end{array} \right\}$ coefficients of f are valid for not too noisy/peaky f ⁸.

⁸F. J. H. and Li. A. Jiménez Rugama (2016). "Reliable Adaptive Cubature Using Digital Sequences". In: *Monte Carlo and Quasi-Monte Carlo Methods: MCQMC, Leuven, Belgium, April 2014*. Ed. by R. Cools and D. Nuyens. Vol. 163. Springer Proceedings in Mathematics and Statistics. arXiv:1410.8615 [math.NA]. Springer-Verlag, Berlin, pp. 367–383; Li. A. Jiménez Rugama and F. J. H. (2016). "Adaptive Multidimensional Integration Based on Rank-1 Lattices". In: *Monte Carlo and Quasi-Monte Carlo Methods: MCQMC, Leuven, Belgium, April 2014*. Ed. by R. Cools and D. Nuyens. Vol. 163. Springer Proceedings in Mathematics and Statistics. arXiv:1411.1966. Springer-Verlag, Berlin, pp. 407–422.



Estimating or Bounding the Error from Simulation Data

Data-driven error bounds based on credible intervals assume that f is an instance of a **Gaussian process** whose parameters are tuned; valid for f that are not outliers. Computation is fast if **covariance kernel** matches the sample⁸.

⁸R. Jagadeeswaran and F. J. H. (2019). “Fast Automatic Bayesian Cubature Using Lattice Sampling”. In: *Stat. Comput.* 29, pp. 1215–1229. doi: 10.1007/s11222-019-09895-9; R. Jagadeeswaran and F. J. H. (2022). “Fast Automatic Bayesian Cubature Using Sobol’ Sampling”. In: *Advances in Modeling and Simulation: Festschrift in Honour of Pierre L’Ecuyer*. Ed. by Z. Botev, A. Keller, C. Lemieux, and B. Tuffin. Springer, Cham, pp. 301–318. doi: 10.1007/978-3-031-10193-9_15.



Problems Involving Several Integrals

$$\text{posterior mean}_j = \frac{\int_{\mathbb{R}^d} x_j \text{likelihood}(\mathbf{x}, \text{data}) \text{prior}(\mathbf{x}) \, d\mathbf{x}}{\int_{\mathbb{R}^d} \text{likelihood}(\mathbf{x}, \text{data}) \text{prior}(\mathbf{x}) \, d\mathbf{x}}, \quad j = 1, \dots, d$$

$$\text{expected output}(\mathbf{s}) = \mathbb{E}[\text{output}(\mathbf{s}, \mathbf{X})] = \int_{\mathbb{R}^d} \text{output}(\mathbf{s}, \mathbf{x}) \varrho(\mathbf{x}) \, d\mathbf{x}, \quad \mathbf{s} \in \Omega$$

$$\text{sensitivity index}_j = \int_{[0,1]^{2d}} f(\mathbf{x}) [f(x_j, \mathbf{z}_{-j}) - f(\mathbf{z})] \, d\mathbf{x} \, d\mathbf{z}, \quad j = 1, \dots, d$$

Stopping criteria have been extended to these cases⁹

⁹F. J. H., Li. A. Jiménez Rugama, and D. Li (2018). “Adaptive Quasi-Monte Carlo Methods for Cubature”. In: *Contemporary Computational Mathematics — a celebration of the 80th birthday of Ian Sloan*. Ed. by J. Dick, F. Y. Kuo, and H. Woźniakowski. Springer-Verlag, pp. 597–619. doi: 10.1007/978-3-319-72456-0; A. G. Sorokin and R. Jagadeeswaran (2023+). “Monte Carlo for Vector Functions of Integrals”. In: *Monte Carlo and Quasi-Monte Carlo Methods: MCQMC, Linz, Austria, July 2022*. Ed. by A. Hinrichs, P. Kritzer, and F. Pillichshammer. Springer Proceedings in Mathematics and Statistics. submitted for publication. Springer, Cham.



Bayesian Logistic Regression for Cancer Survival Data

$\text{logit}(\text{probability of 5 year survival}) = \beta_0 + \beta_1 \text{patient age} + \beta_2 \text{operation year} + \beta_3 \# \text{ positive axillary nodes}$

Logistic regression to estimate β_j from 306 data¹⁰ with error criterion of

absolute error ≤ 0.05 & relative error ≤ 0.5

method	β_0	β_1	β_2	β_3	$\frac{\text{all true}}{\text{all}}$	$\frac{\text{true pos}}{\text{all pos}}$	$\frac{\text{true pos}}{\text{true pos} + \text{false neg}}$
Bayesian & qmcpy	0.0080	-0.0041	0.1299	-0.1569	74%	74%	100%
elastic net	0.0020	-0.0120	0.0342	-0.11478	74%	77%	93%

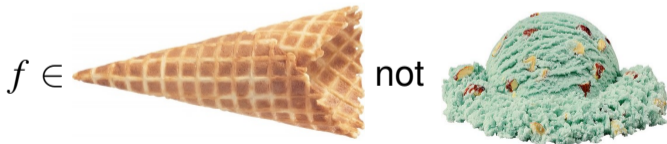
¹⁰SJ Haberman (1976). *Generalized residuals for log-linear models*, proceedings of the 9th International Biometrics Conference.



Lessons from Data-Based Error Bounds

$$\left| \int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x} - \frac{1}{n} \sum_{i=0}^n f(\mathbf{x}_i) \right| \leq \text{some function of } \{(\mathbf{x}_i, f(\mathbf{x}_i))\}_{i=1}^n$$

- Theoretical bounds are **impractical**
- Every data-based error bound can be **fooled**
- There are **better** choices than random replications





Importance Sampling and Control Variates

The original problem may look like

$$\int_{\Omega} g(\mathbf{t}) \, d\mathbf{t}$$

but needs to become $= \int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x}$



Importance Sampling and Control Variates

To facilitate and expedite its solution, one may

- perform a variable transformation, equivalent to **importance sampling**¹¹, and/or
- subtract a trend (**control variate**)¹² with integral zero

$$\int_{\Omega} g(\mathbf{t}) \, d\mathbf{t} \stackrel{\underbrace{=}}{t=\Psi(\mathbf{x})} \int_{[0,1]^d} g(\Psi(\mathbf{x})) \left| \frac{\partial \Psi(\mathbf{x})}{\partial \mathbf{x}} \right| \, d\mathbf{x}, \quad \mathbf{T} = \Psi(\mathbf{X}) \sim \left[\left| \frac{\partial \Psi(\mathbf{x})}{\partial \mathbf{x}} \right|_{\mathbf{x}=\Psi^{-1}(\mathbf{t})} \right]^{-1}$$

$$= \int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x}$$

¹¹A. B. Owen and Y. Zhou (2000). “Safe and Effective Importance Sampling”. In: *J. Amer. Statist. Assoc.* 95, pp. 135–143.

¹²F. J. H., C. Lemieux, and A. B. Owen (2005). “Control Variates for Quasi-Monte Carlo”. In: *Statist. Sci.* 20, pp. 1–31. doi: 10.1214/088342304000000468.



Importance Sampling and Control Variates

To facilitate and expedite its solution, one may

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$$\int_{\Omega} g(\mathbf{t}) \, d\mathbf{t} \stackrel{\underbrace{\quad}}{=} \int_{[0,1]^d} g(\Psi(\mathbf{x})) \left| \frac{\partial \Psi(\mathbf{x})}{\partial \mathbf{x}} \right| \, d\mathbf{x}, \quad \mathbf{T} = \Psi(\mathbf{X}) \sim \left[\left| \frac{\partial \Psi(\mathbf{x})}{\partial \mathbf{x}} \right|_{\mathbf{x}=\Psi^{-1}(\mathbf{t})} \right]^{-1}$$

$$= \int_{[0,1]^d} \left[g(\Psi(\mathbf{x})) \left| \frac{\partial \Psi(\mathbf{x})}{\partial \mathbf{x}} \right| - h(\mathbf{x}) \right] \, d\mathbf{x} = \int_{[0,1]^d} f(\mathbf{x}) \, d\mathbf{x}$$

to make $\text{VAR}(f)$ smaller. Choosing Ψ and h is more art than science.

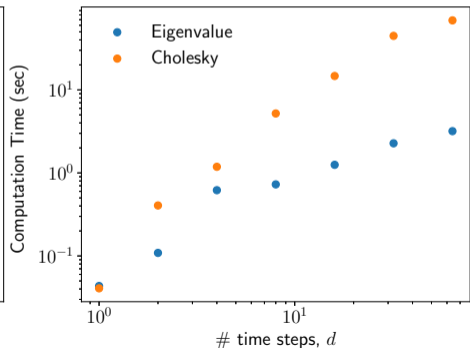
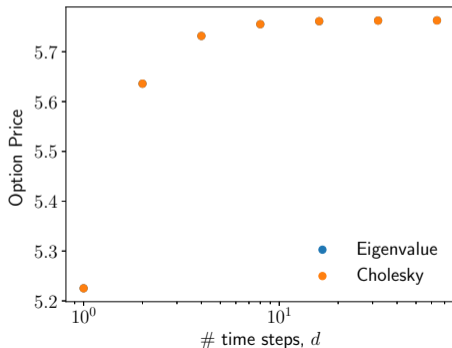
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Asian Option Pricing

- Option payoff is function of a stock price path modeled by a stochastic differential equation
- Option price is expected value of payoff
- True answer is the limit as # of time steps, d , goes to ∞
- Ψ based on the eigenvector-eigenvalue decomposition of the covariance matrix is more efficient than the Cholesky decomposition





Multilevel methods¹³

Large or infinite d when discretizing a continuous stochastic process. Write as

$$\begin{aligned}
\lim_{d \rightarrow \infty} \int_{[0,1]^d} f_d(\mathbf{x}_d) \, d\mathbf{x}_d &= \overbrace{\int_{[0,1]^{d_1}} f_{d_1}(\mathbf{x}_{d_1}) \, d\mathbf{x}_{d_1}}^{\text{low dimensional approximation}} + \overbrace{\int_{[0,1]^{d_2}} [f_{d_2}(\mathbf{x}_{d_2}) - f_{d_1}(\mathbf{x}_{d_1})] \, d\mathbf{x}_{d_2}}^{\text{improvement}} + \dots \\
&+ \int_{[0,1]^{d_L}} [f_{d_L}(\mathbf{x}_{d_L}) - f_{d_{L-1}}(\mathbf{x}_{d_{L-1}})] \, d\mathbf{x}_{d_L} + \lim_{d \rightarrow \infty} \int_{[0,1]^d} f_d(\mathbf{x}_d) \, d\mathbf{x}_d - \int_{[0,1]^{d_L}} f_{d_L}(\mathbf{x}_{d_L}) \, d\mathbf{x}_{d_L} \\
&\approx \underbrace{\frac{1}{n_1} \sum_{i=1}^{n_1} f_{d_1}(\mathbf{x}_{i,d_1}^{(1)})}_{\text{cheap since } d_1 \text{ small}} + \frac{1}{n_2} \sum_{i=1}^{n_2} [f_{d_2}(\mathbf{x}_{i,d_2}^{(2)}) - f_{d_1}(\mathbf{x}_{i,d_1}^{(2)})] + \dots + \underbrace{\frac{1}{n_L} \sum_{i=1}^{n_L} [f_{d_L}(\mathbf{x}_{i,d_L}^{(L)}) - f_{d_{L-1}}(\mathbf{x}_{i,d_{L-1}}^{(L)})]}_{\text{cheap since } n_L \text{ small}}
\end{aligned}$$

where $d_1 < \dots < d_L$ and $n_1 > \dots > n_L$. Evaluation of $f_d(\mathbf{x}_d)$ is typically $\mathcal{O}(d)$.

¹³M. Giles (2013). "Multilevel Monte Carlo methods". In: *Monte Carlo and Quasi-Monte Carlo Methods 2012*. Ed. by J. Dick, F. Y. Kuo, G. W. Peters, and I. H. Sloan. Vol. 65. Springer Proceedings in Mathematics and Statistics. Springer-Verlag, Berlin. doi: 10.1007/978-3-642-41095-6. 19/25



Message

Problems in Bayesian inference, uncertainty quantification, quantitative finance, (high energy) physics, etc. require **numerically** evaluating

$$\underbrace{\mathbb{E}[f(\mathbf{X})]}_{\text{expectation}} = \underbrace{\int_{\Omega} f(\mathbf{x}) \varrho(\mathbf{x}) \, d\mathbf{x}}_{\text{integral}}, \quad \mathbf{X} \sim \varrho$$

There is value in





- **low discrepancy** sampling, aka **quasi-Monte Carlo** methods
- data-driven **error bounds**
- **flattening** and **reducing the effective dimension** of the integrand
- quality quasi-Monte Carlo **software** like **qmcpy**
- physicists and quasi-Monte Carlo theorists **collaborating** more

Slides at speakerdeck.com/fjhickernell/argonne2023maytalk

Jupyter notebook with computations and figures [here](#). Visit us at qmcpy.org



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