# Practical Statistics <br> for Particle Physics 



Daniel Whiteson, UC Irvine HCPSS, 2014: Lecture 2

## Outline

I. Mathematical preliminaries II. Fitting III. Data models IV. Hypothesis testing
V. Tools and examples

## Models

# Full MC Fast MC Effective models Data-driven models 

## Uncerfainties

We have a recipe for

## $f($ data | theory)

But is it right?

## Uncertainties

We have a recipe for

## $f($ data | theory)

But is it right?

Theory has lots
of nuisance parameters: cross-sections, LO, NLO... showering details hadronization details detector response

There is some point in NP space which gives the most accurate model
but we don't know where it is!

## Systematics

## The Good

NP can be constrained in some control region. Uncertainty decreases with luminosity.
eg. Background cross-section
B-tagging efficiency
Jet energy scale
The Ugly
Underlying theoretical approach eg PYTHIA vs HERWIG

The Bad
Parameters of underlying heuristic eg PYTHIA tunes

(Pekka Servino)

## Example



## Example



## Example



## Extrapolation




Measure background in one region, extrapolate knowledge to another. How well do we know rates of jet production, tails of MET?

## Extrapolation



This is

$$
\begin{aligned}
& \text { his is } \\
& \text { cleverness happens! } \\
& \text { hap main one }
\end{aligned}
$$ rapolate knowledge Mother. How well do we know rates of jet production, tails of MET?

## Jeł energy scale




Total uncertainty
Absolute scale
-Relative scale

- Pile-up NPV=8
$\square$ Jet flavor
${ }^{\nabla}$ Time stability
Anti-k $\mathrm{R}=\mathbf{0 . 5} \mathrm{PF}$
$\eta_{\text {jet }} \mathrm{l}=0$

Many steps in jet production Lots of opportunities for mistakes Calibrate in ii, photon+jet Extrapolate to your dataset


## the shift method



## Generating

 samples at arbitrary values of NP can be expensive!Often, just generate a few and interpolate.

## Histogram interpolation

A.L. Read / Nuclear Instruments and Methods in Physics Research A 425 (1999) 357-360


## dimensions

systematic $A$


## dimensions

$$
\begin{array}{cc|c|c} 
& \begin{array}{c}
\text { systematic A }
\end{array} & \begin{array}{c}
\text { We often } \\
\text { generate }
\end{array} \\
\text { these off-axis } \\
\text { points. }
\end{array}
$$

## dimensions

$$
\begin{aligned}
& \text { We almost }
\end{aligned}
$$

> We rely on
> linear interpolation

## Uncertainties

Scale variation


## Uncertainty:

 shift renormalization, factorization scales$$
\text { by } 2,1 / 2
$$

measure change.

Why 2, 1/2?
Just convention

Not 1 sigma!

## Generałors



fast-MC model

## fast MC model



## Begin with generated events

 but rather than simulating microphysics, smear particles according to resolution.
## Delphes

fast simulation


## fast MC model

Less accurate, but same issues as full MC: No analytic PDF Uncertainties in simulation


## The dream

## f(data | final-state particles P)

$x \mathrm{f}($ final state particles $\mathrm{P} \mid$ showered particles S$)$
$\times \mathrm{f}($ showered particles $\mathrm{S} \mid$ hard scatter products M)
$\times f($ hard scatter products $M \mid$ theory $)$

## Sum over all possible intermediate P,S,M

## ME approach

## If we have a parametrized detector response,

 can we parameterizef(data|final-state particles P)
$\times \mathrm{f}($ final state particles $\mathrm{P} \mid$ showered particles S$)$
x $f($ showered particles $S$ | hard scatter products $M)$
$\times f($ hard scatter products $M \mid$ theory $)$

## ME approach

## Yes we can!

$$
P\left(\mathbf{x} \mid M_{t}\right)=\frac{1}{N} \int d \Phi\left|\mathcal{M}_{t t}\left(p ; M_{t}\right)\right|^{2} \prod_{j e t s} f\left(p_{i}, j_{i}\right) f_{P D F}\left(q_{1}\right) f_{P D F}\left(q_{2}\right)
$$

| Phase-space <br> Integral |
| :--- |



Matrix Element


## ME approach

$$
\begin{aligned}
P\left(\mathbf{x} \mid M_{t}\right) & =\frac{1}{N} \int d \Phi\left|\mathcal{M}_{t \bar{t}}\left(p ; M_{t}\right)\right|^{2} \prod_{j e t s} f\left(p_{i}, j_{i}\right) f_{P D F}\left(q_{1}\right) f_{P D F}\left(q_{2}\right) \\
& \\
\begin{array}{ll}
\text { Phase-space } \\
\text { Integral }
\end{array} & \begin{array}{l}
\begin{array}{l}
\text { Matrix } \\
\text { Element }
\end{array}
\end{array}
\end{aligned}
$$

Transfer functions reflect a very complex process
By necessity, approximations, and therefore uncertainties.

## Data-driven model

## Example: dark matter



## Mono-jeł

## Missing --................. q/g <br> Momentum

## Event display



CMS Experiment at LHC, CERN
Data recorded: Sun Oct 30 16:05:09 2011 CEST
Run/Event: 180250 / 878954337
Lumi section: 481

## Backgrounds



Final state:
jet + MET

Process:
$Z \rightarrow v v$, with jet

## Backgrounds

## How to estimate?

Idea: $Z \rightarrow v \vee$ from $Z \rightarrow$ II


Approach:
(1) measure Z to II + je†
(2) scale by known branching ratios


## Defails

$$
N[Z(v v)]=N[Z(I I)] \times B F[Z(v v)] / B F[Z(I I)]
$$

## Defails

$$
\begin{aligned}
N[Z(\mathrm{vv})]= & N[Z(I I)] \times B F[Z(\mathrm{vv})] / B F[Z(I I)] \\
& N[Z(I I)]=N(I I)-N(b g) / \varepsilon
\end{aligned}
$$

## Defails

$$
N[Z(\mathrm{vv})]=\mathrm{N}[\mathrm{Z}(\mathrm{II})] \times \mathrm{BF}[\mathrm{Z}(\mathrm{vv})] / \mathrm{BF}[\mathrm{Z}(\mathrm{II})]
$$

$$
N[Z(I I)]=N(I I)-N(b g) / \varepsilon
$$



CMS PAS EXO-12-048

## Effective Model

## Effective Model



## High-level arguments

dependence of background eg "smooth background"

## shape of signal

eg "narrow resonance"

## Effective Model



## Summary of models

## MC simulation

- sample of events from on/off simulation
- estimate PDF from events

Fast MC simulation

- simpler generation model
- still estimate PDF from events

Data-driven model

- extrapolate from control regions

Effective model

- parametrized functional form


## Summary of models

|  | Pros | Cons |
| :---: | :---: | :---: |
| MC Simulation | $\begin{array}{c}\text { detailed descr } \\ \text { of } \text { micro physics }\end{array}$ | $\begin{array}{c}\text { very slow } \\ \text { must reconstruct PDF }\end{array}$ |

Data-driven Calculations by Nature
Effective model
fast,
fast
Fast MC physical justification
approximate
Extrapolations from CR have uncertainties approximate no details of underlying effects

## Hypothesis Testing

## Hypothesis Testing

Claim Discovery

No Claim of Discovery

BSM Particle BSM Particle is real is not real

| True | False <br> Positive <br> Positive <br> Type I error |
| :---: | :---: |

False
Negative
Type II error
$\beta$, power=1- $\beta$

True
Negative

## Example



A threshold makes sense.
Choice of position balances
Type I/II errors

Typically:
fix $\alpha$ minimize $\beta$

## Generalize

Hypothesis Testing

## HO H1

Parameter
Estimation
H 1
cross-section

## More complicated



## Test statistic

Reduce vector of observables to 1 number


How to build distribution of TS? (Usually MC) How to choose TS?
(K. Cranmer)

## Neyman-Pearson

## Statement of the problem:

Given some prob that we wrongly reject the Null hypothesis

$$
\alpha=P\left(x \notin W \mid H_{0}\right)
$$

Find the region $W$ (where we accept $H_{0}$ ) such that we minimize the prob

$$
\beta=P\left(x \in W \mid H_{1}\right)
$$

|  | BSM Particle <br> is real | BSM Particle <br> is not real |
| :---: | :---: | :---: |
| Claim <br> Discovery | True <br> Positive | False <br> Positive |
| No Claim <br> Type I error |  |  |
| False |  |  | | Negative |
| :---: | :---: |
| Type II error |
| $\beta$, power=1- $\beta$ |$\quad$| True |
| :---: |
| Negative |

## Neyman-Pearson

## NP lemma says that the best

 test statistic is the likelihood ratio:$$
\frac{P\left(x \mid H_{1}\right)}{P\left(x \mid H_{0}\right)}>k_{\alpha}
$$

(Gives smallest $\beta$ for fixed $\alpha$ )

|  | BSM Particle <br> is real | BSM Particle <br> is not real |
| :---: | :---: | :---: |
| Claim <br> Discovery | True <br> Positive <br> Fo Claim <br> N Discovery <br> Negative | False <br> Positive |
| Type I error |  |  |
| Type II error |  |  |
| $\beta$, power=1- $\beta$ |  |  |$\quad$| True |
| :---: |
| Negative |

## What does the TS do?

Finds a region in variable space

(K. Cranmer)

## How to find NP

## Isolate some

 feature in which two theoriesSM, SM+X can be best distinguished.

Standard Model
SM+X

- Collider Data

The data can tell us which hypothesis is preferred via a likelihood ratio:

$$
\frac{L_{S M+X}}{L_{S M}} \quad \frac{P(\text { data } \mid S M+X)}{P(\text { datata } \mid S M)}
$$

## e.g.



## But...

## Reality is more complicated.

The full space can be very high dimensional.

Calculating likelihood in d-dimensional space requires $\sim 100^{\mathrm{d}} \mathrm{MC}$ events.

Standard Model

feature 1

## ML tools



## Neural Nełworks

Essentially a functional fit with many parameters


## Function

Each neuron's output is a function of the weighted sum of inputs.

## Goal

find set of weights which give most useful function

## Learning

give examples, back-propagate

## Neural Nełworks

Essentially a functional fit with many parameters


## Problem:

Networks with > 1 layer are very difficult to train.

## Consequence:

Networks are not good at learning non-linear functions. (like invariant masses!)

## In short:

Can'† just throw 4-vectors at NN.

## Search for Inpuł

## ATLAS-CONF-2013-108

## Can't just use 4v

Can't give it too many inputs

## Painstaking search through input feature space.

| Variable | VBF |  |  | Boosted |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\tau_{\text {lep }} \tau_{\text {lep }}$ | $\tau_{\text {lep }} \tau_{\text {had }}$ | $\tau_{\text {had }} \tau_{\text {had }}$ | $\tau_{\text {lep }} \tau_{\text {lep }}$ | $\tau_{\text {lep }} \tau_{\text {had }}$ | $\tau_{\text {had }} \tau_{\text {had }}$ |
| $m_{r T}^{\text {MMC }}$ | - | - | - | - | - | - |
| $\Delta R(\tau, \tau)$ | $\bullet$ | - | - |  | $\bullet$ | - |
| $\Delta \eta\left(j_{1}, j_{2}\right)$ | $\bullet$ | $\bullet$ | - |  |  |  |
| $m_{j 1, h_{2}}$ | - | - | $\bullet$ |  |  |  |
| $\eta_{i_{1}} \times \eta_{j_{2}}$ |  | - | - |  |  |  |
| $p_{\text {T }}^{\text {Iotal }}$ |  | - | - |  |  |  |
| sum $p_{\text {T }}$ |  |  |  |  | - | $\bullet$ |
| $p_{\mathrm{T}}\left(\tau_{1}\right) / p_{\mathrm{T}}\left(\tau_{2}\right)$ |  |  |  |  | - | - |
| $E_{T}^{\text {miss }} \phi$ centrality |  | - | - | - | - | - |
| $x_{\tau 1}$ and $x_{\tau 2}$ |  |  |  |  |  | - |
| $m_{\tau \tau, j 1}$ |  |  |  | - |  |  |
| $m_{\ell_{1}, \ell_{2}}$ |  |  |  | - |  |  |
| $\Delta \phi_{\ell_{1}, \ell_{2}}$ |  |  |  | - |  |  |
| sphericity |  |  |  | - |  |  |
| $p_{T}^{\ell_{1}}$ |  |  |  | - |  |  |
| $p_{\mathrm{T}}^{j_{1}}$ |  |  |  | - |  |  |
| $E_{\mathrm{T}}^{\mathrm{miss}} / p_{\mathrm{T}}^{\ell_{2}}$ |  |  |  | - |  |  |
| $m_{\mathrm{T}}$ |  | - |  |  | - |  |
| $\min \left(\Delta \eta_{\ell_{1} \ell_{2} \text { jets }}\right)$ | - |  |  |  |  |  |
| $j_{3} \eta$ centrality | - |  |  |  |  |  |
| $\ell_{1} \times \ell_{2} \eta$ centrality | $\bullet$ |  |  |  |  |  |
| $\ell \eta$ centrality |  | - |  |  |  |  |
| $\tau_{1,2} \eta$ centrality |  |  | - |  |  |  |

Table 3: Discriminating variables used for each channel and category. The filled circles identify which variables are used in each decay mode. Note that variables such as $\Delta R(\tau, \tau)$ are defined either between the two leptons, between the lepton and $\tau_{\text {had }}$, or between the two $\tau_{\text {had }}$ candidates, depending on the decay mode.

## Search for Input

## ATLAS-CONF-2013-108

## Can't just use 4v

| Variable | $\tau_{\text {lep }} \tau_{\text {lep }}$ |
| :---: | :---: |
| $\frac{m_{r T}^{M M C}}{D^{\prime}}$ | VBF |

Can't give it +
many ing AlsO trues,
etc $\vdots$
Painstakins $B D T S$, through inp
feature space.


Table 3: Discriminating variables used for each channel and category. The filled circles identify which variables are used in each decay mode. Note that variables such as $\Delta R(\tau, \tau)$ are defined either between the two leptons, between the lepton and $\tau_{\text {had }}$, or between the two $\tau_{\text {had }}$ candidates, depending on the decay mode.

## Deep nełworks



## New tools

 let us train deep networks.How well do they work?

## Real world applications


(a)

(e)

(b)

(f)

(c)

(g)

(d)

(h)

Head turn: DeepFace uses a 3-D model to rotate faces, virtually, so that they face the camera. Image (a) shows the original image, and (g) shows the final, corrected version.

## Paper

Deep Learning in High-Energy Physics: Improving the Search for Exotic Particles

P. Baldi, ${ }^{1}$ P. Sadowski, ${ }^{1}$ and D. Whiteson ${ }^{2}$<br>${ }^{1}$ Dept. of Computer Science, UC Irvine, Irvine, CA 92617<br>${ }^{2}$ Dept. of Physics and Astronomy, UC Irvine, Irvine, CA 92617

## arXiv: 1402.4735 <br> Accepted in Nature Comm.

## What does the TS do?

Finds a region in variable space

(K. Cranmer)

## Test statistic

At LEP, this was used:
Define $\mu$ to be signal strength, $\mu=0$ is no signal $\mu=1$ is theory prediction

$$
Q_{L E P}=L_{s+b}(\mu=1) / L_{b}(\mu=0)
$$

Where the nuisance parameters are fixed to their nominal values

## Test statistic

Define $\mu$ to be signal strength, $\mu=0$ is no signal
At LEP, this was used: $\mu=1$ is theory prediction

$$
Q_{L E P}=\frac{L(\text { data } \mid \mu=1, b, \nu)}{L(\text { data } \mid \mu=0, b, \nu)}
$$

This also means the background estimate doesn't vary.

## Tevałron

Still consider two points $(0,1)$ but now float the NPs at those points

$$
Q_{T E V}=L_{s+b}(\mu=1, \hat{\hat{\nu}}) / L_{b}\left(\mu=0, \hat{\hat{\nu}}^{\prime}\right)
$$

the model is adapted to the data even in the signal region

## LHC

## Profile likelihood

$$
\lambda(\mu=0)=\frac{L(\text { data } \mid \mu=0, \hat{b}(\mu=0), \hat{\hat{v}}(\mu=0))}{L(\operatorname{data} \mid \hat{\mu}, \hat{b}, \hat{v})}
$$

fit best value of NPs at $\mu=0$ and at best fit value of $\mu$

## Two fits to data

$$
\begin{aligned}
& \lambda(\mu=0)=\frac{L(\text { data } \mid \mu=0, \hat{\hat{b}}(\mu=0), \hat{\hat{v}}(\mu=0))}{L(\text { data } \mid \hat{\mu}, \hat{b}, \hat{v})} \quad L(\text { data } \mid \mu=0, \hat{\hat{b}}, \hat{\hat{\nu}})
\end{aligned}
$$




## $p$ values



$$
p_{0}=P\left(q_{0} \geq q_{o}^{o b s}\right)
$$

## $p$ values


> $\mathrm{P}_{\mu}=$ probability to observe data or less signal-like under signal+b hypothesis

## Philosophy

Bayesian
\&
Frequentist

## Bayesian

## Data: fixed

Parameter values: unknown Probability: our lack of knowledge
PDFs over parameters: sensible

## Frequentist

Data: one example from ens. Parameter values: fixed (even if unknown) Probability: rate of occurance
PDFs over parameters: not sensible

## Bayesian Prob.

Bayes theorem:

$$
\begin{aligned}
& P(A \mid B)=\frac{P(A \cap B)}{P(B)} \\
& P(B \mid A)=\frac{P(A \cap B)}{P(A)}
\end{aligned}
$$

rearrange:

$$
\begin{gathered}
P(A \mid B) P(B)=P(A \cap B)=P(B \mid A) P(A) \\
P(A \mid B)=\frac{P(B \mid A) P(A)}{P(B)}
\end{gathered}
$$

## In Picłures

## P, Conditional P, and Derivation of Bayes' Theorem

 in Pictures

$$
\begin{aligned}
& \mathbf{P}(\mathbf{A})=\frac{0}{\square} \\
& \mathbf{P}(\mathbf{B})=\frac{0}{\square} \\
& \mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{0}{\square} \\
& \mathbf{P}(\mathbf{A} \cap \mathbf{B} \mid \mathbf{B})=\frac{0}{\square}
\end{aligned}
$$

$$
\mathbf{P}(\mathbf{A}) \times \mathbf{P}(\mathbf{B} \mid \mathbf{A})=\frac{0}{\square} \times \frac{0}{0}=\frac{0}{\square}=\mathbf{P}(\mathbf{A} \cap \mathbf{B})
$$

$$
\mathbf{P}(\mathbf{B}) \times \mathbf{P}(\mathbf{A} \mid \mathbf{B})=\frac{0}{\square} \times \frac{0}{\square}=\frac{0}{\square}=\mathbf{P}(\mathbf{A} \cap \mathbf{B})
$$

$$
\Rightarrow P(B \mid A)=P(A \mid B) \times P(B) / P(A)
$$

## Example 1

$\mathrm{P}($ data |theory) ! $=\mathrm{P}($ theory $\mid$ data)
Theory = (male or female)
Data $=$ (pregnant | not pregnant)
P(pregnant | female) ~ 3\%

## BUT

P(female | pregnant) >99\%

## Example 2

## Higgs search

Expected bg = 0.1
Expected signal $=10$
$P(N \mid$ no Higgs $)=0.1$
$P(N \mid$ Higgs $)=10.1$
What is $\mathrm{P}($ Higgs $\mid \mathrm{N}=8)$ ? $\quad P(H \mid N=8)=\frac{P(N=8 \mid H) P(H)}{P(N=8)}$
Depends on $\mathrm{P}(\mathrm{H})$ !
(K Cranmer)

## Parameter estimation

## Bayesian parameter estimation:

Want to know the probability that some parameter $\theta$ is in some range $\left[\theta_{0}, \theta_{1}\right]$

- Or -

Want to find a range $\left[\theta_{0}, \theta_{1}\right]$ that has probability of 0.95

## Parameter estimation

## Bayesian parameter estimatio-

Want to know the probat"
parameter $\theta$ is in som

- or -

Want ' of

## How?

The probability that the true value is inside an interal is:

$$
1-\alpha=\int_{\theta_{l o}}^{\theta_{h i}} p(\theta \mid x) d \theta
$$

For lower or upper limits, choose zero or infinity as boundaries. where we integrate out the nuisance parameters:

$$
p(\theta \mid x)=\int d \nu p(\theta, \nu \mid x)
$$

where

$$
p(\theta, \nu \mid x)=\frac{p(x \mid \theta, \nu) p(\theta, \nu)}{p(x)}
$$

These integrals can be very hard to do if the space is high dimensional.

## Priors

## Choice of prior $\mathrm{p}(\theta)$

- important but subjective choice



## Priors

## Choice of prior $\mathrm{p}(\theta)$

- Example: measuring Higgs cross-section
- Want to be unbiased: choose uniform prior?

$$
\sigma=[0, \wedge] \rightarrow P=k
$$

- But $\sigma$ and mass relationship makes this prior not flat in mass

-no uninformative prior across all transformations

